

ARTIFICIAL INTELLIGENCE, ALGORITHMIC RECOMMENDATIONS AND COMPETITION

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We develop a framework to assess the economic impact of algorithmic recommendations, conceptualized as creating *personalized prominence* in consumer search. Our approach integrates realistic algorithms trained on synthetic data with a flexible model of consumer preferences and product differentiation. To demonstrate the utility of our framework, we apply it to three widely debated issues. First, we evaluate how algorithmic recommendations influence market concentration and the diversity of consumer choices. Second, we analyze their effects on equilibrium prices and consumer welfare. Finally, we explore the potential for platforms to manipulate recommendations to prioritize more profitable products.

1. INTRODUCTION

Recommender systems (RSs) are AI-based algorithms that employ user feedback—such as ratings, clicks, or other activities—to predict users’ potential interest in items they have not yet experienced. Digital platforms, which serve as intermediaries between users and suppliers of these items, routinely collect this feedback and use algorithmic predictions to provide personalized recommendations, helping users navigate the vast array of available options. These recommendations already exert a substantial influence on consumer demand (Jannach and Juvovac, 2019), and their influence is expected to grow further with advancements in AI technology and the ongoing accumulation of user data by platforms.¹

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¹While this paper specifically focuses on the recommendation of products to consumers, RSs have applications beyond this context. For example, social media platforms use RSs to curate users’ feeds, financial institutions leverage them for investor products (e.g., Robo-advisors), employment agencies utilize them to match workers with employers, and academia applies RSs to tasks such as identifying suitable referees for editors or suggesting relevant articles to scholars.

Despite their importance, the economic analysis of RSs remains underdeveloped, primarily due to the challenges of modeling them in a meaningful yet tractable way. Most existing analyses assume that consumers simply follow algorithmic recommendations. In contrast, this paper proposes a novel framework that conceptualizes RSs as creating *personalized prominence* in consumer search. This framework acknowledges that consumers do not always adhere to platform recommendations. Instead, these recommendations offer personalized pre-search information — allowing consumers to start their search with products that are expected to align well with their preferences while maintaining the flexibility to continue searching if the suggested products turn out to be poor matches.

This perspective requires embedding the analysis of algorithmic recommendations within a fully developed model of consumer search. However, we cannot rely on the standard search model, which assumes entirely idiosyncratic preferences. Our focus is on collaborative-filtering recommendation systems—arguably the most popular and widely used class of recommendation algorithms—which function by identifying patterns of similarity and dissimilarity among users and products. If consumer preferences were entirely idiosyncratic, any observed correlations in the data would be coincidental, rendering recommendations ineffective or even misleading. To ensure consistency, we must therefore adopt a model of consumer preferences and product characteristics that exhibits systematic similarities and differences among products and consumers, allowing the algorithms to generate valuable predictions.

And yet, these predictions cannot be perfect, as in reality, platforms observe feedback for only a tiny fraction of all consumer-product pairs, and this feedback may provide only noisy signals of the true value of products to consumers. As a result, algorithmic estimates are inherently subject to small-sample biases.

Integrating these key elements results in a highly complex framework that challenges the application of analytical methods. This paper proposes a numerical approach to identifying and analyzing the impact of RSs within controlled, albeit synthetic, economic settings. Building on the computer science literature, state-of-the-art algorithms are implemented and trained using synthetic data generated by a model of consumer preferences and product differentiation. This approach ensures control over both the quality and quantity of information provided to the algorithm. Consumers’ optimal search patterns and firms’ optimal pricing strategies are then determined numerically, contrasting the scenario in which consumers receive personalized recommendations with a benchmark where they rely solely on unassisted search. To address concerns about external validity, an extensive robustness analysis is conducted.

To demonstrate the potential applications of this novel approach, which integrates RSs into a search framework, we revisit three issues widely debated among scholars and policymakers. First, we analyze the impact of algorithmic recommendations on market concentration and the diversity of consumer choices. The central question is whether RSs help consumers discover niche products that would otherwise go unnoticed or whether they create a rich-get-richer dynamic, disproportionately promoting a few popular items and leading to more highly concentrated markets.² This analysis specifically focuses on subscription-based platforms such as Netflix and Spotify, where users pay a fixed fee and do not incur additional costs for each item consumed.

Turning to platforms like Amazon, where each product is individually priced, we secondly examine the effect of RSs on equilibrium prices and consumer welfare. Algorithmic recommendations can influence prices even in the absence of price discrimination by altering

²Similar issues also arise in other settings. For example, it has been contended that the use of RSs may cause political polarization in social media and a loss of diversity in culture. See, for instance [Abdollahpouri and Mansoury \(2020\)](#), [Abdollahpouri et al. \(2021\)](#).

how consumers search for products, thereby affecting consumer demand. Furthermore, algorithmic recommendations directly impact consumer welfare by changing the level of equilibrium search and improving the quality of matches between consumers and the products they ultimately purchase.

Finally, we investigate the possibility that platforms may manipulate their recommendations to favor more profitable products at the expense of genuinely superior ones. Often, these favored products are those supplied by the platform itself, which is why the practice is commonly referred to as *self-preferencing*. Existing analyses of self-preferencing, which assume that consumers mechanically follow the platform's recommendations, may overestimate the impact of this practice.

We now provide a brief preview of our substantive results, deferring a more detailed discussion to the subsequent sections. First, we find that RSs tend to favor mass-market products over niche ones. This is due to a "uniformity" effect, where the algorithm overestimates consumer similarity, leading it to recommend products that align too frequently with the preferences of the median consumer. This bias disappears only when the quantity and quality of data available to the algorithm are very high, far beyond what is typically observed in practice.

Second, we find that RSs lead firms to raise prices, even when price discrimination is excluded by assumption.³ However, compared to a scenario without algorithmic recommendations, RSs also improve the matching between consumers and products, reducing the need for costly search. While we generally observe a positive total effect, a decline in consumer surplus may occur in markets with relatively few products, predominantly horizontal product differentiation, and high search costs.

Even when consumers benefit from RSs compared to their absence, we also demonstrate that as algorithms gain access to more and better data, consumer surplus may eventually decrease. In other words, the relationship between information and consumer welfare may follow an inverted-U curve. Initially, consumer surplus rises as the algorithm's information level increases, allowing consumers to more easily find products that match their preferences. However, higher levels of information also drive up prices. Beyond a certain threshold, this negative effect dominates, and further increases in information lead to a reduction in consumer surplus. This pattern holds regardless of the source of information variation, whether due to increased data quantity, noise reduction, or other factors.

Finally, we find that when platforms manipulate their recommendations, the prices of over-recommended products tend to decrease. This mitigates the negative welfare impact of such practices and limits their profitability, suggesting that the profit-maximizing level of manipulation may be relatively small.

We discuss more fully the relationships between each set of results and the relevant literatures in the subsequent sections, where we present the results in greater detail. Methodologically, the two contributions most closely related to this paper are [Lee and Wright \(2021\)](#) and [Castellini et al. \(2023\)](#). However, these papers do not embed the analysis of RSs within a framework of individual search⁴ and do not investigate the impact of RSs on product market competition.⁵

³[Lee and Musolf \(2023\)](#) empirically uncover price effects of algorithmic recommendations that align with the findings of our analysis.

⁴[Lee and Wright \(2021\)](#) assess the information value of RS algorithms by comparing them to purely random choices, whereas [Castellini et al. \(2023\)](#) use complete information as their benchmark.

⁵More broadly, this paper contributes to the recent literature on the implications of AI for industrial organization. Most of this literature has, so far, focused specifically on the issue of algorithmic collusion: see, for instance, [Calvano et al. \(2020\)](#), [Johnson et al. \(2023\)](#), [Asker et al. \(2023\)](#), and [Klein \(2021\)](#). More recent work has added empirical evidence ([Assad et al., forthcoming](#)) and theoretical results ([Possnig, 2023](#), [Banchio and Mantegazza, 2023](#)).

The rest of the paper is divided into two parts. In the first part, we present our analytical framework. Section 2 offers a self-contained introduction to the latent-factor, collaborative-filtering algorithms utilized. Section 3 presents a model of product differentiation that exhibits systematic similarities and differences among consumers and products. It is a model of the “address” variety that, for consistency, mirrors the latent-factor structure of the algorithms. The model encompasses different combinations of horizontal and vertical product differentiation, as well as both “niche” and “mass” products. Section 4 describes the data the algorithms use and how they are trained. Section 5 introduces the search theoretic framework that we use to model consumer behavior. Finally, Section 6 specifies the baseline parameterization of the model in preparation for the numerical analysis.

The second part of the paper presents our substantive findings. Section 7 examines the impact of RSs on market concentration, while Section 8 analyzes the effect of RSs on equilibrium prices. In Section 9, we analyze the consequences of varying the amount and quality of information, demonstrating an inverted-U relationship between information and consumer welfare. Section 10 analyzes the case where the platform manipulates its recommendations. In the concluding section, we discuss potential extensions to our work and the policy implications of our findings. An online appendix includes details that are omitted from the main text. Extensions of narrower interest are included in supplementary material [Calvano et al. \(2025\)](#).⁶

2. MODEL-BASED RECOMMENDER SYSTEMS

We focus on latent-factor, collaborative-filtering systems, a class of algorithms that includes the winner of the Netflix Prize.⁷ According to ([Aggarwal, 2016](#), p. 91), these algorithms are “considered to be state-of-the art in recommender systems.” [Rokach et al., 2022](#) suggest that they are widely used in practice.⁸

2.1. *The basic problem*

Consider a finite set of users $i = 1, 2, \dots, I$, a finite set of items $j = 1, 2, \dots, J$ and a sparse matrix $\tilde{\mathbf{R}}$ of size $I \times J$ which is commonly referred to as the *rating* matrix in the computer science literature (we adopt this terminology throughout). Existing entries \tilde{r}_{ij} represent the rating of user i for item j when such rating is either reported or inferred from user behavior, and the entries for the unobserved user-item pairs are missing. (Economists naturally interpret these ratings as indicative of the utility user i derived from item j .) The objective of the algorithm is to estimate the full matrix of true ratings, denoted as \mathbf{R} .

Collaborative filters re-estimate the observed values and fill in the missing ones exploiting the correlation structure of observed ratings. The underlying assumption is that users whose observed ratings are similar are likely to have similar ratings for unobserved items, while items whose ratings are similar across observed users are likely to be rated similarly by unobserved users.

⁶This supplementary material is available at the DOI link: <https://doi.org/10.7910/DVN/EREVKI>

⁷In 2006, Netflix launched a one million dollar prize for the first RS that could improve the performance of Cinematch, their algorithm for predicting ratings, by at least 10%. The challenge provided participants with a rich data set. After three years, the prize was awarded in 2009 to BellKor’s Pragmatic Chaos team for improving Netflix’s algorithm by 10.06%. A new challenge in 2010 was discontinued due to concerns that the anonymity of the data provided had been breached. For more details, see [Koren et al. \(2009\)](#).

⁸The main advantage of latent factor methods based on matrix factorization is their high parallelizability, enabling efficient management of massive, sparse, and dynamically changing datasets. This capability makes them ideal for delivering a large volume of recommendations, especially in challenging environments where items frequently change or quickly become outdated.

2.2. Latent factors

The latent-factor, collaborative filtering algorithms we focus on explain the observed user-item ratings and predict the unobserved ones by using hidden (latent) factors that represent inherent properties of both items and users. This approach reduces the dimensionality of the rating matrix by representing the ratings in terms of a relatively small number of variables.

Analytically, the true rating of item j by user i is viewed as the inner product of a vector of user-specific parameters \mathbf{t}_i and a vector of item-specific parameters \mathbf{v}_j :

$$r_{ij} = \sum_{h=1}^H t_{ih} v_{jh}, \quad (1)$$

where H is the number of latent factors. In matrix notation, $\mathbf{R} = \mathbf{T}\mathbf{V}'$, where \mathbf{T} and \mathbf{V} are the $(I \times H)$ and $(J \times H)$ matrices formed by the I vectors \mathbf{t}_i and the J vectors \mathbf{v}_j , respectively.

The variable t_{ih} may be thought of as user i 's proclivity for factor h , and the variable v_{jh} as product j 's affinity to factor h . In applications, the H factors may have a semantic interpretation. For the algorithm, however, they need not have any specific meaning.

2.3. Estimation procedure

The algorithm estimates the parameters $\hat{\mathbf{T}}$ and $\hat{\mathbf{V}}$ by minimizing some measure of the distance between the estimated ratings,

$$\hat{r}_{ij} := \sum_{h=1}^H \hat{t}_{ih} \hat{v}_{jh} \quad (2)$$

and the observed ones, \tilde{r}_{ij} . Using the Euclidean distance, the estimated parameters $\hat{\mathbf{T}}$ and $\hat{\mathbf{V}}$ solve:

$$\min_{\hat{\mathbf{T}}, \hat{\mathbf{V}}} \sum_{(i,j) \in S} \left(\tilde{r}_{i,j} - \sum_{h=1}^H \hat{t}_{ih} \hat{v}_{jh} \right)^2, \quad (3)$$

where S denotes the set of all pairs (i, j) for which r_{ij} is (imperfectly) observed.⁹ Using the completed matrix $\hat{\mathbf{R}} = \hat{\mathbf{T}}\hat{\mathbf{V}}'$, the algorithm can generate rankings of all items for all users, allowing it to provide personalized recommendations.

It is important to note that in collaborative-filtering RSs, attributes and tastes are estimated *jointly*, which marks a departure from much of the structural empirical work on demand estimation based on the Random Utility Model. Unlike in those estimations, where it is typically assumed that the choice attributes (i.e., the v_{jh} s) are observable while preferences (i.e., the t_{jh} s) are not, RSs observe neither product attributes nor consumer tastes and estimate both.¹⁰

In real-life applications, the density d of the matrix $\hat{\mathbf{R}}$ of observed rating is quite low, often below 2%.¹¹ Consequently, estimates obtained from such limited data may exhibit small sample

⁹Problem (3) assumes that the algorithm uses a pre-specified number of latent factors H , as we do in our baseline analysis. However, in a robustness check, we allow the algorithm to estimate the value of H using an internal cross-validation routine, as described in the online appendix. The results remain nearly unchanged.

¹⁰The case where some attributes are observable by the RS can be handled by using hybrid RSs (also known as ensemble RSs). By symmetry, these algorithms can also deal with the case where consumer tastes are partly observable. These possible extensions are left for future work.

¹¹The *density* of a sparse matrix is the fraction of non-empty cells, and the *sparsity* is the fraction of empty cells. Therefore, $\text{density} = 1 - \text{sparsity}$.

biases. In particular, the estimation procedure (3) may cause overfitting. A common approach for addressing this problem is to use regularization techniques. The idea is to penalize large values of the coefficients in the matrices $\hat{\mathbf{T}}$ and $\hat{\mathbf{V}}$ so as to decrease the variance of $\hat{\mathbf{R}}$. The actual minimization problem then becomes:

$$\min_{\hat{\mathbf{T}}, \hat{\mathbf{V}}} \sum_{ij \in S} \left(\tilde{r}_{i,j} - \sum_{h=1}^H \hat{t}_{ih} \hat{v}_{jh} \right)^2 + \lambda_t \sum_{i=1}^I \sum_{h=1}^H \hat{t}_{ih}^2 + \lambda_v \sum_{j=1}^J \sum_{h=1}^H \hat{v}_{jh}^2, \quad (4)$$

where the non-negative parameters λ_t and λ_v are the regularization weights. The weights are chosen by a cross-validation procedure described in the online appendix.¹²

3. PRODUCTS AND PREFERENCES

In this section, we present a model of consumer preferences and product characteristics that mirrors the latent-factor structure employed by the algorithms. This ensures that the estimation can be based on a correctly specified model of the economic environment, and in cases of mis-specification, it allows us to control its nature and extent. The model is rich enough to accommodate the coexistence of niche and mass products, as well as consumers with both common and more eccentric preferences.

3.1. A latent-factor model of product differentiation

We examine a monopolistic platform that serves as an intermediary between I buyers (indexed by $i = 1, 2, \dots, I$) and the sellers of J products (indexed by $j = 1, 2, \dots, J$). Consumer i 's willingness to pay for product j , in monetary units, consists of a deterministic component \bar{u}_{ij} that is specific to each product-consumer pair, along with an idiosyncratic shock ε_{ij} :

$$u_{ij} = \bar{u}_{ij} + \varepsilon_{ij}. \quad (5)$$

The shocks ε_{ij} are normally i.i.d. with zero mean and variance σ_ε^2 . (They serve to smooth out the demand functions, guaranteeing the existence of a price equilibrium in pure strategies.) The systematic component, on the other hand, is assumed to be:

$$\bar{u}_{ij} = \sum_{h=1}^H t_{ih} v_{jh}. \quad (6)$$

Note the analogy with the structure of the ratings r_{ij} that the algorithm tries to predict. From an economic viewpoint, H represents the number of characteristics of the products that are valued by consumers (akin to different dimensions of product quality), $v_{jh} \geq 0$ the level of the h -th characteristic in product j , and $t_{ih} \geq 0$ the value that consumer i attaches to it. It therefore appears that this latent-factor formulation is consistent with the Lancasterian approach, which posits that consumers derive "satisfaction from characteristics that [...] cannot be purchased directly, but are incorporated in goods" (Lancaster et al., 1974, p. 567).

We assume that the sum $\sum_{h=1}^H t_{ih}^2$ is constant across consumers, so all consumers have the same total willingness to pay for quality. The interpretation is that consumers may be heterogeneous in tastes but not in income. This assumption makes the model trivial if $H = 1$; therefore, we focus on the case $H \geq 2$. With no further loss of generality, we can normalize the total willingness to pay to one: $\sum_{h=1}^H t_{ih}^2 = 1$.

¹²The appendix also shows that our results are robust to the choice of these weights.

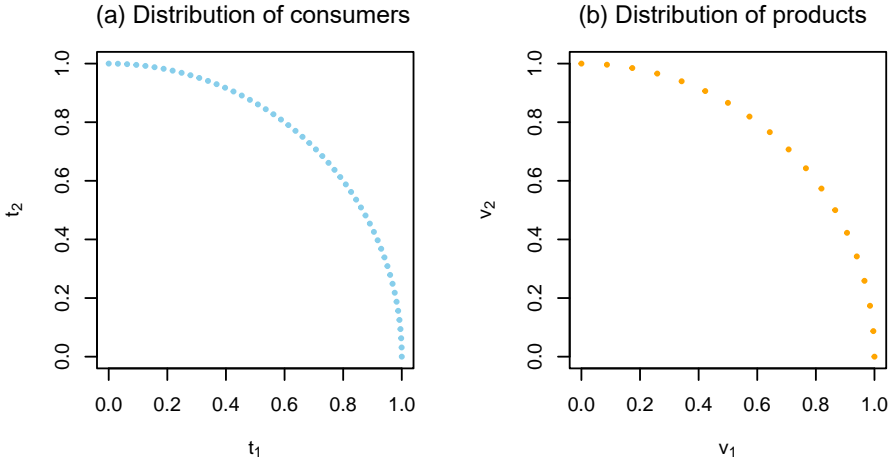


FIGURE 1.—The distribution of consumers and products in the two-factor model with purely horizontal differentiation.

3.1.1. Horizontal differentiation

The case where product differentiation is purely horizontal is obtained when all goods have the same “total” quality $\sum_{h=1}^H v_{jh}^2$, which can then also be normalized: $\sum_{h=1}^H v_{jh}^2 = 1$. In this case, both products and consumers can be represented as points on the portion of the unit hyper-sphere that lies in the non-negative orthant.¹³ Specifically, we assume that both the J products and the I consumers are equally spaced (see Figure 1).

For each possible consumer type (i.e., for each vector $\mathbf{t} = (t_1, \dots, t_H)$), there is a different ideal product $\mathbf{v} = (v_1, \dots, v_H)$, which under horizontal differentiation is $\mathbf{v} = \mathbf{t}$. When consumers purchase a product that differs from their ideal product, they incur a “transportation cost” akin to that considered in the standard Hotelling model. For example, when $H = 2$, using polar coordinates $\{t_{i1} = \cos \theta, t_{i2} = \sin \theta\}$ and $\{v_{j1} = \cos v, v_{j2} = \sin v\}$, it is easy to see that the transportation cost is $1 - \cos(\theta - v)$.¹⁴

The restriction to the non-negative orthant implies that products are not symmetric even if product differentiation is purely horizontal. To see this, consider for instance the case of three products. Clearly, the *central product*, located on the 45-degree line, enjoys higher demand compared to the two peripheral ones. Thus, it qualifies as a natural candidate for becoming a “mass” product. Conversely, the *peripheral products*, which are located on the x - and y -axes in Figure 1, represent the “niche” products. The presence of both mass and niche products enables us to study whether RSs lead to a “superstar effect” or a “long-tail effect.” The issue will be taken up in section 7.

¹³With two factors, the model becomes similar to Wolinsky (1983), with the twist that products and consumers are restricted to the non-negative portion of the unit circle.

¹⁴Thus, the transportation cost is a convex function of the distance between products and consumers. See the online appendix for more details.

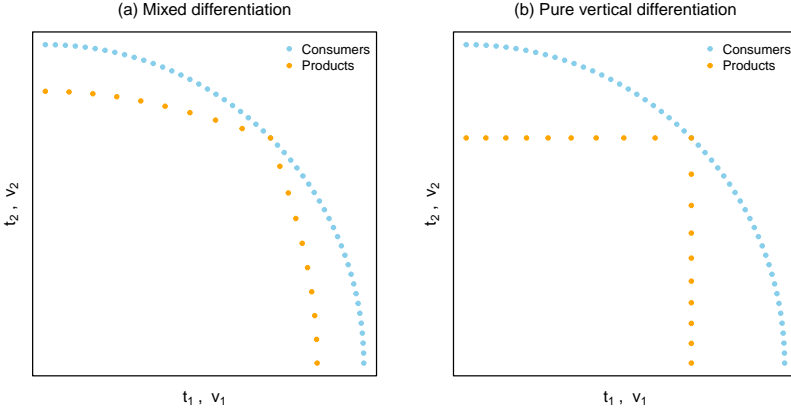


FIGURE 2.—The distribution of products under vertical, or mixed, differentiation.

3.1.2. Vertical differentiation

The assumption that product differentiation is purely horizontal can be relaxed, by allowing for the possibility that the quality index $\sum_{h=1}^H v_{jh}^2$ may vary across products. This extension introduces vertical differentiation into the model.

For example, one could assume that products are located on the square rather than on the circle (Figure 2, right panel). In this case, all consumers agree on their favorite product, which lies on the 45-degree line.

Intermediate cases between purely horizontal and purely vertical differentiation can be obtained by varying the shape of the product locus from a circle to a square. This can be done parsimoniously, by means of a single parameter α that ranges from $\alpha = 0$ (circle, horizontal differentiation) to $\alpha = 1$ (square, vertical differentiation). To be precise, for each product \mathbf{v}_j on the circle, define its radial projection on the square as $\mathbf{v}'_j = \frac{\mathbf{v}_j}{\sqrt{2} \max_h [v_{jh}]}$. By taking the convex combination $\alpha \mathbf{v}_j + (1 - \alpha) \mathbf{v}'_j$ of the circle and the square, one can then generate all intermediate combinations of vertical and horizontal differentiation.

As soon as $\alpha > 0$, a bunch of consumers (i.e., those close to the 45-degree line) will prefer the central product. This is an additional reason why the central product stands out as a candidate mass product.¹⁵

4. DATA

After purchasing a product, consumers may explicitly or implicitly report a rating to the platform. These ratings are the data used by the algorithms to estimate the match values.

In the baseline scenario, we assume that the reported rating \tilde{u}_{ij} is equal to the utility actually experienced, u_{ij} , plus a noise term:

$$\tilde{u}_{ij} = u_{ij} + \epsilon_{ij}. \quad (7)$$

¹⁵In principle, the flexibility of the model permits the consideration of scenarios where the highest-quality product does not necessarily coincide with the central product. Geometrically, this can be achieved by employing a rectangle instead of a square. Additionally, it is also possible to explore situations where peripheral products possess higher quality. However, these possibilities are not pursued here.

The reporting noise ϵ_{ij} is normally i.i.d. with zero mean and variance σ_ϵ^2 . The higher the variance σ_ϵ^2 , the less informative the ratings reported by consumers.

Later on, we will consider extensions where the report is made on a Likert scale. It is important to note that ratings on a two-level Likert scale may be available to a platform even without explicit reporting by users. For instance, the platform could monitor whether a consumer has purchased an inspected product, or watched a movie until the end.

As noted, RSs typically have access to information about only a small fraction of all consumer-product pairs. What determines which ratings are observed and which are not? We consider two alternative data generation processes, depending on whether the data have been generated based on recommendations provided by the algorithm itself in the past or not.

4.1. Randomly generated data

In our baseline analysis, we assume that the algorithm itself does not play a role in generating the data it uses. Specifically, we assume that for each user, the platform observes ratings for a fixed number of products, drawn randomly and independently across users. The fraction of consumer-product pairs for which ratings are observed corresponds to the density d of the matrix $\tilde{\mathbf{R}}$ introduced in Section 2.

4.2. Endogenous data

Alternatively, we consider the case where the ratings observed by the platform have been created based on recommendations provided by the algorithm itself in the past. In this case, the matrix $\tilde{\mathbf{R}}$ is filled in gradually, in a number of successive steps indexed by $\tau = 0, 1, 2, \dots$. Initially, $\tilde{\mathbf{R}}_0$ is the empty matrix; that is, a matrix with all entries missing. At step $\tau = 1$, each consumer tries a randomly drawn product resulting in ratings \tilde{u}_{ij} for the RS. The matrix $\tilde{\mathbf{R}}_1$ therefore has a density of $1/J$.

At each subsequent step, the RS uses $\tilde{\mathbf{R}}_\tau$ to produce an estimate $\hat{\mathbf{R}}_\tau$ of the true matrix according to the estimation procedure described in Section 2 and recommends to each consumer i the product with the highest rating:

$$j^*(i, \tau) = \arg \max_j \hat{r}_{ij\tau}. \quad (8)$$

This product is then tried by the consumer,¹⁶ and the corresponding rating reported to the platform. Thus, $\tilde{\mathbf{R}}_{\tau+1}$ is obtained by adding to $\tilde{\mathbf{R}}_\tau$ one and only one entry $\tilde{u}_{ij^*(i, \tau)}$ for each of the I rows (i.e. for each consumer). If $\tilde{u}_{ij^*(i, \tau)}$ is not missing in $\tilde{\mathbf{R}}_\tau$, its value is overwritten. All other elements of $\tilde{\mathbf{R}}_\tau$ are unchanged.

Thus, $\tilde{\mathbf{R}}_{\tau+1}$ is basically equal to $\tilde{\mathbf{R}}_\tau$ plus one new entry per consumer, which could be either a previous value overwritten or a brand new one. As this process unfolds, the density of $\tilde{\mathbf{R}}_\tau$ increases. The algorithm's learning phase ends when the matrix $\tilde{\mathbf{R}}$ achieves a pre-specified

¹⁶Note that during the learning phase, for simplicity, we abstract from the possibility that consumers may disregard the recommendation and engage in their own search. However, this possibility plays a crucial role in our analysis of the effects of recommendations once the learning phase has been completed, as we will see below.

density d .¹⁷ At this juncture, we analyze the effects of the recommendations ultimately generated by the algorithm.

Data endogeneity creates a feedback loop extensively discussed in the computer science literature: the algorithm gathers better information on the user-item pairs it has selected in the past. This is often referred to as the “bias in the algorithms” (even though a more precise designation might be “bias in the data”). This bias emerges because, even if the final density matches that of random data, the user-item pairs observed by the algorithm are somewhat correlated, diminishing the quality of the available information. Consequently, the precision of estimates and the quality of recommendations decline.

5. SEARCH

We now integrate the above model into a search framework. This allows us to conceptualize recommendations as generating a form of personalized prominence. In this context, consumers do not necessarily follow platform recommendations but can initiate their search with products specifically recommended to them, rather than with randomly chosen ones.

5.1. Product markets

Before proceeding, it is important to note that although the willingness to pay for any two of the J products is correlated to some extent, not all products traded on the platform necessarily serve as direct substitutes for one another.

For example, the platform may offer both movies and books. Individuals who enjoy watching war movies may also like historical novels, while those who prefer wildlife documentaries may lean towards science books. The algorithm benefits from pooling all products together to leverage the correlation in consumers’ evaluations. Nevertheless, books and movies belong to distinct markets.

To account for this, we partition the set of all J products into distinct subsets (markets) \mathcal{M} and assume that each consumer buys at most one product from each market \mathcal{M} . We denote the number of products belonging to market \mathcal{M} as $m_{\mathcal{M}}$.

5.2. Prior information

In general, consumer behavior in a market is shaped by the information individuals have regarding product availability, prices, and match values. Given that the specific role of RSs is to estimate the match values, we evaluate their impact by *holding other information constant*.

Different assumptions can be made regarding this other information. In the classic search model of [Wolinsky \(1986\)](#) and [Anderson and Renault \(1999\)](#), consumers are aware of product availability but do not know individual prices or match values; they are only aware of the distribution of product characteristics and equilibrium prices. Acquiring information on individual prices and match values requires costly search. However, more recent literature on search has also explored an alternative scenario, where consumers know product prices but lack knowledge of match values.¹⁸

¹⁷In practice, achieving a target density level may require assuming that the algorithm operates in *exploration mode* for a certain fraction of periods, recommending products randomly. Without this, the process may reach a steady state with a density lower than the target. In our simulations, the case of endogenous data corresponds to the lowest possible probability of being in exploration mode, typically around 10 percent. By increasing the probability of exploration mode, we can generate data that lie between purely exogenous and endogenous.

¹⁸Yet another possibility is that consumers may not be fully aware of product availability. However, the analysis of [Aridor et al. \(2022\)](#) suggests that this factor may be of lesser relevance.

It is important to note that, to focus on the specific function of RSs (which is to estimate match values), in the first scenario, the platform’s recommendations should be based solely on the estimated match values. In the second scenario, recommendations should also account for product prices. Without this distinction, the analysis would conflate the effects of match value information (which is the focus of our analysis) with those of price information.

For simplicity, we focus on the model of [Wolinsky \(1986\)](#) and [Anderson and Renault \(1999\)](#) in this paper. The advantage of this approach is that, in the absence of recommendations, consumer search would be purely random. In contrast, when consumers have information on prices, they would engage in directed search. This complicates the analysis, so we defer further investigation into this case to future research.

5.3. Random search benchmark

Based on his prior information and knowledge of his own preferences, each consumer i calculates the probability distribution, denoted as $F_{i,\mathcal{M}}(s_i)$, for his expected surplus $s_{ij} = u_{ij} - p_j$ from the products $j \in \mathcal{M}$, where p_j denotes the price.

In the benchmark without recommendations, consumer i initially assigns an equal expected surplus to all products, which is the unconditional mean of s_i based on the probability distribution $F_{i,\mathcal{M}}(s_i)$. Following this, he has the option to inspect the products, incurring a unit search cost of c_s . Lacking specific knowledge about individual products, consumers must search randomly. When they sample a product, they observe its price and match value.

To ensure stationarity, we assume sampling with replacement and no recall.¹⁹ Under this assumption, consumers will continue to search until they reach a cut-off level of surplus, denoted as $\hat{s}_{i,\mathcal{M}}$, which is the solution to:

$$\int_x^\infty (s_i - x) dF_{i,\mathcal{M}}(s_i) = c_s. \quad (9)$$

Intuitively, at the optimum, the expected benefit of one additional search must be equal to its cost.

5.4. Personalized prominence

Compared to this benchmark, we investigate a scenario where platforms utilize algorithms to estimate match values and provide personalized recommendations based on these estimates. As soon as the recommendations are just minimally informative, consumers will prioritize examining the products suggested by the RS first. This introduces a form of prominence, as described by [Armstrong et al. \(2009\)](#). Here, however, different products may be prominent for different consumers, so prominence is personalized.

¹⁹Under the alternative assumption of sampling with replacement and perfect recall, since consumers know that the products are equally spaced, the probability distribution $F_{i,\mathcal{M}}(s_i)$ changes upon inspecting a product and observing its location, implying that the cutoffs vary after each visit. Since the calculation of the cutoffs is by far the most time-consuming part of the numerical analysis, the property of stationarity is analytically convenient. One could dispense with the assumption of replacement without losing stationarity by assuming that products are distributed randomly instead of being equally spaced. However, this would blur the distinction between mass and niche products. In any case, the difference with the more common assumption of sampling with replacement tends to vanish as the number of products grows large and disappears with infinitely many products.

Denote the product j with the highest estimated match value for consumer i in market \mathcal{M} as $j^*(i, \mathcal{M})$.²⁰ If the platform does not strategically manipulate the recommendations, it will suggest product $j^*(i, \mathcal{M})$ to consumer i .²¹ Upon receiving the recommendation, consumer i inspects the recommended product $j^*(i, \mathcal{M})$, observes its price $p_{j^*(i, \mathcal{M})}$, and obtains an assessment of the match value, $u_{ij^*(i, \mathcal{M})}$. At this point, he has two options: either purchase the recommended product or conduct further search, at the same unit cost c_s as in the benchmark. Any additional search conducted is random, similar to the case without RSs. Thus, the consumer will choose to continue searching if the expected surplus $u_{ij^*(i, \mathcal{M})} - p_{j^*(i, \mathcal{M})}$ falls below the cut-off $\hat{s}_{i, \mathcal{M}}$.

6. BASELINE PARAMETRIZATION

In this section, we specify the baseline parameterization of the model in preparation for the numerical analysis.

We choose the baseline values of the parameters so that the model matches some key properties of the Netflix Challenge (see footnote 6 and Table I). In the dataset provided by Netflix to participants in the challenge, the ratio between the number of users and items was approximately 30, so we set $\frac{I}{J} = 30$. Likewise, the density d of the matrix of ratings $\tilde{\mathbf{R}}$ is set to the same value as in Netflix Challenge, i.e., 1.2%.²²

To reduce the computational burden of the simulations, we scale down the number of products and consumers by a factor of approximately 20 compared to the Netflix problem. Thus, we set $J = 801$ and hence $I = 24,030$. (The number of products is odd so that their distribution can be symmetric around the central product.)

	Environment				Recommender System			
	Users I	Items J	Observations	Density	H	ℓ	H	ℓ
Netflix	500,000	17,000	100,000,000	1.2%	40	4.84	100	1.93
Baseline	24,030	801	231,265	1.2%	2	4.65	5	1.86

TABLE I: The baseline scenario and the Netflix Challenge

The winner of the Netflix Challenge used a number of latent factors H of either 40 or 100. (The trade-off is that more latent factors allow for more flexibility but increase the number of variables to be estimated with the same observations, reducing the precision of the estimates.) To calibrate the corresponding values for our baseline model, we match the ratio between the number of observations and the number of parameters to be estimated,

$$\ell = \frac{d}{H} \frac{J \times I}{(J + I)}. \quad (10)$$

²⁰As noted, we assume that recommendations are solely based on the match value to maintain consistency with the unassisted search benchmark, in which consumers are unaware of product prices.

²¹The extension where the platform recommends a list of products is left for future work. If the platform manipulates the recommendations, all that changes is that the recommended product may not be $j^*(i, \mathcal{M})$.

²²Computer scientists use several benchmark datasets to evaluate RSs. Some of these datasets have densities similar to Netflix's, such as the MovieLens 10M Dataset, which contains users' movie ratings collected by the GroupLens research project at the University of Minnesota and has a density of 1.34%. However, other benchmark datasets have lower densities. In the robustness analysis, we investigate the sensitivity of our findings to variations in density.

This leads us to consider two possible values of H , namely $H = 2$ (which is our baseline choice and corresponds to $H = 40$ in the Netflix Challenge) and $H = 5$ (which is taken up among the extensions and corresponds to $H = 100$).²³

For several other parameters of the model, the Netflix challenge or similar datasets provide no guidance. We proceed as follows. First, we consider a grid of values of the number of (equally spaced) products in the market, $m_{\mathcal{M}}$, and the degree of vertical differentiation, α , with $m_{\mathcal{M}}$ ranging from 7 to 31 with a step size of 6 and α ranging from 0 (pure horizontal differentiation) to 1 (pure vertical differentiation), with a step size of 0.1.²⁴

Second, we explore different values of the unit search cost parameter c_s . To provide an intuition for their practical relevance, we relate c_s to the fraction of consumers who purchase the recommended product without further search. We let this fraction vary between 70% and 95%. This translates into a range of values for c_s from 0.002 to 0.006.

Finally, remember that our baseline model assumes that consumers report the utility level $\tilde{u}_{ij} = \bar{u}_{ij} + \varepsilon_{ij} + \epsilon_{ij}$, inclusive of the normally distributed shocks. Both shocks have zero mean, so it remains to specify their variance. In the baseline analysis, the standard deviation of the taste shocks σ_ε is set equal, for each consumer i , to 10% of the standard deviation of the distribution of \bar{u}_{ij} across the J products. This value is about as low as it can be without prejudicing convergence of the iterative procedure for the calculation of equilibrium prices.²⁵ As for the reporting noise, in the baseline scenario we set $\sigma_\epsilon = 2\sigma_\varepsilon$.

For each set of parameters, we conduct 100 simulations with different realizations of the uncertainties regarding the utility shocks (ε_{ij}), the reporting noise (ϵ_{ij}), and the products included in $\tilde{\mathbf{R}}$. The results are then averaged across sessions.

In the main text, for clarity, we present only the case where $m_{\mathcal{M}} = 19$ and $c_s = 0.004$, which corresponds to around 85% of consumers not searching further,²⁶ and we use a step size of 0.25 for α . The online appendix provides detailed results for the entire grid of α , $m_{\mathcal{M}}$, and c_s values.

After analyzing the baseline scenario, we conducted several robustness checks. Since most of these involve changes to the quality or quantity of information available to the algorithm, a more detailed description is provided in Section 9, where we examine the impact of information on consumer welfare. For now, it suffices to say that the results for the baseline scenario, as reported in the following sections, remain generally robust across all these extensions. In the main text, we highlight only the most significant variations, while a more comprehensive presentation of the robustness analysis is available in the online appendix and supplementary material.

²³One problem that arises in this extension is the absence of a general solution for evenly positioning points along the section of the unit hyper-sphere located in the positive orthant when $H > 3$. To overcome this difficulty, we have developed an iterative algorithm inspired by the k -means method to approximate the precise positioning of products. More details are provided in the supplementary material.

²⁴Note that while our assumption that consumers at the outset have no clue of product quality is reasonable when $\alpha = 0$, as α increases, it becomes progressively less tenable. Consequently, while we have examined all conceivable combinations of horizontal and vertical product differentiation, our analysis particularly pertains to scenarios where vertical differentiation is limited.

²⁵When σ_ε is lower, the iterative procedure may become trapped in cycles, suggesting that equilibria may involve mixed strategies.

²⁶Although this fraction is high, it serves to better highlight the effects of RSs. Intuitively, as the search cost decreases, the influence of recommendations tends to diminish. In the extreme case where the search cost becomes negligible, consumers will choose to inspect all products before making a purchase, rendering the impact of recommendations negligible as well.

7. CONCENTRATION OF CHOICES

In this section, we concentrate on subscription platforms like Netflix or Spotify, where consumers pay a fixed fee and can access any product without additional charges. Thus, the prices of individual products, p_j , are all zero. For this scenario, the primary question we address is the impact of RSs on market concentration.

We specifically revisit two opposing views of RSs that have emerged in the debate over the effects of RSs: the “long-tail” and “superstar” views. As described by [Bar-Isaac et al. \(2012\)](#), the former refers to a scenario where RSs increase the popularity of niche products, while the latter refers to a scenario in which they increase the popularity of mass products. The long-tail view was first articulated by [Anderson \(2008\)](#), who argues that online markets exhibit a long-tail phenomenon due to the larger selection of products available to suppliers and consumers easier access to niche products.²⁷ In contrast, [Fleder and Hosanagar \(2009\)](#), among others, have found that RSs may reinforce the popularity of already popular items, resulting in a decrease in diversity at the aggregate level, even though individual-level diversity may increase. (Note that these views are not mutually exclusive: in principle, when there are more than three products, both effects can coexist, at the expense of intermediate products.)

Here, we revisit this debate using a well-structured model of consumer preferences and product characteristics and explicitly modeling consumer search. This approach enables us to precisely identify the effect of deploying RSs on the concentration of users’ choices.²⁸

7.1. *Quality of recommendations*

Before proceeding, we assess the effectiveness of RS in matching consumers with products. This evaluation is based on the welfare obtained by consumers under the assumption of zero prices, which is maintained in this section. [Table II](#) compares the average utility, net of search costs, obtained by the average consumer across the 100 simulations under RS and individual search, with the standard deviation shown in parenthesis.

As further benchmarks, the table also reports the net utility that the average consumer would obtain under complete information (this is less than 1 because not all consumers t can find a perfect match $v = t$, or because, when $\alpha > 0$, the quality of the ideal product may be less than 1) or when choices are entirely random.

The results show that the RS generates a significant utility gain, despite having limited information. Both the magnitude of the gain and its sources depend on the degree of vertical differentiation in the market. The gain is higher, the more vertically differentiated the products. As for the sources, the total gain can be decomposed into better matching and reduced search. With purely horizontal differentiation ($\alpha = 0$), it turns out that two-thirds of the gain comes from lower total search costs. However, when α is larger, a substantial portion of the gain is due to the RS’s ability to identify the best match.

²⁷[Brynjolfsson et al. \(2011\)](#) claim that the long-tail phenomenon persists even when holding product availability constant.

²⁸The existing literature in marketing and computer science often posits assumptions at odds with economic consumer theory, motivated by applications of RSs other than consumer demand.

α	0	0.25	0.5	0.75	1
Full information	0.9957	0.9510	0.9199	0.9020	0.8963
Random choice	0.7969	0.7563	0.7158	0.6753	0.6347
<i>Unassisted search</i>	0.9694 (0.0000)	0.9225 (0.0000)	0.8844 (0.0000)	0.8569 (0.0000)	0.8411 (0.0000)
<i>Recommender System</i>	0.9834 (0.0001)	0.9396 (0.0001)	0.9102 (0.0001)	0.8944 (0.0001)	0.8911 (0.0001)
$\frac{RS-U_n}{U_n} \times 100$	1.44% (0.01%)	1.85% (0.01%)	2.92% (0.01%)	4.37% (0.01%)	5.95% (0.01%)

TABLE II: Consumer average utility, net of search costs (SD in parentheses).

7.2. Tails

We now investigate whether RSs lead to an increase in the market share of niche products, resulting in a long-tail effect. To explore this question, we examine the average market shares of the most peripheral products, i.e., the ones lying on the x - and y -axis in Figure 1. As noted, these products are natural candidates for being niche products in our framework. The results are presented in Table III.

α	0	0.25	0.5	0.75	1
<i>Unassisted search</i>	0.0325 (0.0000)	0.0263 (0.0000)	0.0189 (0.0000)	0.0103 (0.0000)	0.0029 (0.0000)
<i>Recommender System</i>	0.0258 (0.0012)	0.0216 (0.0010)	0.0158 (0.0007)	0.0091 (0.0004)	0.0018 (0.0000)
$\frac{RS-U_n}{U_n} \times 100$	-20.66% (3.82%)	-18.16% (3.79%)	-16.39% (3.79%)	-11.63% (3.41%)	-38.11% (1.73%)

TABLE III: Market share of niche products (i.e., the two most peripheral products in Figure 1).

Contrary to the long-tail hypothesis, the market shares of niche products are consistently lower with RSs than in the benchmark. (In fact, we observe a decrease in the market shares not only of the two most extreme products, but also of the nearby ones.) The decline ranges from one-tenth to more than a third. Similar results are found for any number of products and any level of the search cost, as well as in all the extensions discussed in Section 9. Therefore, we conclude that our findings do not support the notion that RSs encourage the diffusion of niche products.

7.3. Superstars

The alternative view of RSs in marketing and computer science is that they tend to produce superstars. In order to determine whether there is any evidence of a superstar effect, we perform the same analysis on the central product, which is located on the 45-degree line in Figure 1.

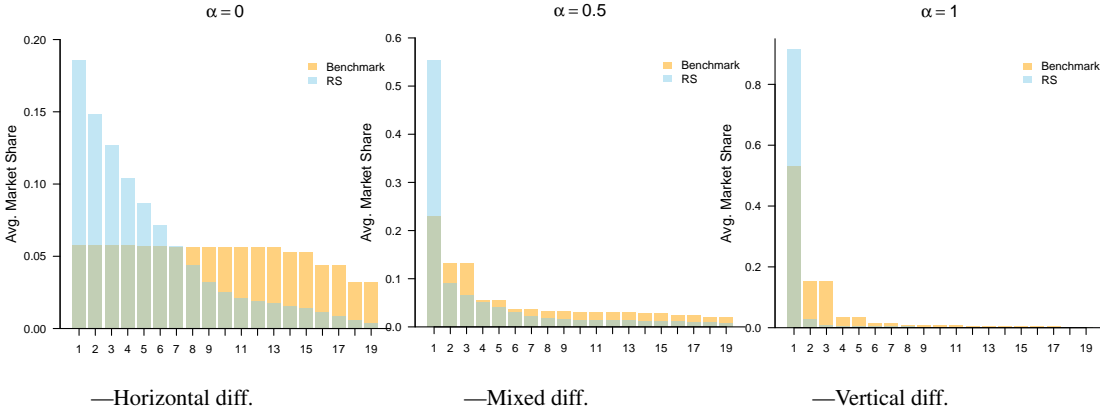


FIGURE 3.—Avg Market share by product rank

α	0	0.25	0.5	0.75	1
<i>Unassisted search</i>	0.0565 (0.0000)	0.1291 (0.0000)	0.2302 (0.0000)	0.3650 (0.0000)	0.5312 (0.0000)
<i>Recommender System</i>	0.0682 (0.0074)	0.3091 (0.0096)	0.5539 (0.0065)	0.7545 (0.0047)	0.9153 (0.0034)
$\frac{RS-U_n}{U_n} \times 100$	20.61 (13.10)	139.41 (7.40)	140.67 (2.82)	106.73 (1.29)	72.30 (0.64)

TABLE IV: Market share of the central product.

Remember that the central product stands out due to two reasons: its central position, which places it closest to the median consumer, and its higher total quality when $\alpha > 0$. As shown in Table IV, with RSs, the market share of the central product increases significantly. Again, this result remains true varying the number of products and the level of the search cost, as well as in all the extensions considered in Section 9.

The rise in the market share of the central product is smaller when product differentiation is purely horizontal ($\alpha = 0$),²⁹ but in fact, the RS has a strong tendency to create superstars even in this case. This tendency is not fully apparent in Table IV because with $m_{\mathcal{M}} = 19$, several products are “central” to some extent, and the ones selected by the RS as superstars vary randomly from session to session.

To control for this effect, we ranked products based on their market share and calculated the average share of the most popular product, the second most popular, and so forth. The resulting distribution is illustrated in Figure 3, for both the RS scenario and the benchmark of unassisted search. It appears that the RS creates a substantial superstar effect even when $\alpha = 0$, with the most popular product earning a market share more than three times larger than in the benchmark.

²⁹The rise is also less pronounced when product differentiation is predominantly vertical. This is because, in such cases, consumers can readily identify the superior product to a significant degree even without relying on the RS.

7.4. The degree of market concentration

The combination of superstar effect and reverse long-tail effect results in a significant increase in market concentration. To quantify this phenomenon, we use the Herfindahl-Hirschman Index (HHI), as shown in Table V.³⁰ The increase in the HHI is substantial for all types of product differentiation, but is particularly significant for intermediate values of α .

For large values of α , the increase in market concentration may be seen as a benign outcome, as it indicates that the algorithm is able to effectively identify and promote the superior product to consumers. However, for low values of α , it is more concerning. In this case, the RS appears to create its own “champions” without any clear objective basis for favoring one product over others.

α	0	0.25	0.5	0.75	1
<i>Unassisted search</i>	538.5979 (0.0000)	650.1194 (0.0000)	1,057.3305 (0.0000)	1,902.0176 (0.0000)	3,323.1593 (0.0000)
<i>Recommender System</i>	1,078.0114 (11.4180)	1,593.3218 (37.1896)	3,329.5506 (57.4467)	5,788.5495 (61.6039)	8,410.1934 (56.8680)
$\frac{RS-U_n}{U_n} \times 100$	100.15% (2.12%)	145.08% (5.72%)	214.90% (5.43%)	204.34% (3.24%)	153.08% (1.71%)

TABLE V: The Herfindahl-Hirschman Index of market concentration.

7.5. Estimation biases and the “uniformity effect”

What, then, is the source of the superstar effect and the increase in market concentration? Our analytical framework provides a valuable tool for addressing this question. By ensuring the correct model specification by the RS, we can attribute these effects to estimation biases resulting from small sample sizes. To explore this issue further, we examine how the estimates $\hat{\mathbf{T}}$ and $\hat{\mathbf{V}}$ compare with the true coefficients \mathbf{T} and \mathbf{V} . Given that the increase in market concentration is particularly concerning in the case of purely horizontal product differentiation ($\alpha = 0$), we focus our analysis on this scenario.

To gain some preliminary insight, we start by randomly selecting one session. For this session, Figure 4 depicts the algorithm’s estimates of both consumers’ and products’ latent vectors, $\{(\hat{t}_{i1}, \hat{t}_{i2})\}_{i=1}^I$ and $\{(\hat{v}_{j1}, \hat{v}_{j2})\}_{j=1}^J$. The red dashed curve represents the “true” consumers and products, as in Figure 1, while the orange dots represent the “virtual” consumers \hat{t}_i estimated by the algorithm, and the blue dots the “virtual” products \hat{v}_j . The size of the disk around each blue dot indicates how frequently that product is recommended.

Note that the virtual consumers $(\hat{t}_{i1}, \hat{t}_{i2})$ are fully described by the ratio $\frac{\hat{t}_{i2}}{\hat{t}_{i1}}$, with the level of the \hat{t} s being irrelevant. This means that equi-proportional changes in t_1 and t_2 will not affect consumers’ choices as long as the market is covered. On the other hand, for the virtual products $(\hat{v}_{j1}, \hat{v}_{j2})$, both the ratio $\frac{\hat{v}_{j2}}{\hat{v}_{j1}}$ and the level of the \hat{v} s matter. The former represents the product’s estimated “type,” the latter the estimated “quality.”

³⁰We follow the common practice in industrial organization of normalizing the HHI so that it ranges between 0 and 10,000. We have also considered other indices of market concentration, such as the Gini index or the fraction of products that carry a positive market share, with similar results.

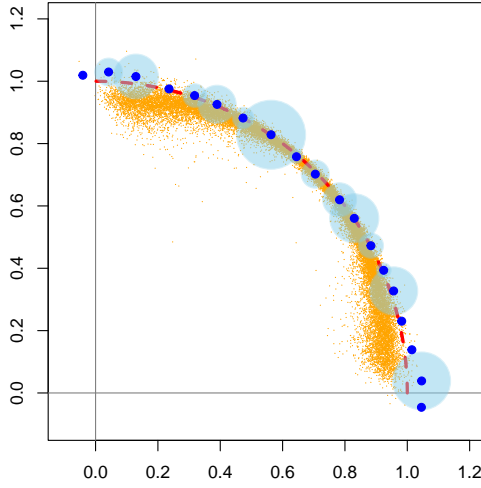


FIGURE 4.—The estimated location of consumers (orange dots) and products (blue dots), plotted against their true location (the red dashed curve), for a single session with $\alpha = 0$. The size of the disk around the blue dots represents the product's market share.

In principle, therefore, the algorithm can make three types of errors: (i) the estimated consumer tastes may differ from the true ones, resulting in $\frac{\hat{t}_{i2}}{\hat{t}_{i1}} \neq \frac{t_{i2}}{t_{i1}}$; (ii) the estimated product types may differ from the true ones, resulting in $\frac{\hat{v}_{j2}}{\hat{v}_{j1}} \neq \frac{v_{j2}}{v_{j1}}$; and (iii) the estimated product qualities may differ from the true ones, resulting in $\sum_{h=1}^2 \hat{v}_{jh}^2 \neq \sum_{h=1}^2 v_{jh}^2 (= 1)$.

The session shown in Figure 4 illustrates all three types of estimation biases. Firstly, the orange dots are clustered together towards the 45-degree line, indicating that the algorithm tends to overestimate consumer uniformity. We shall refer to this tendency as the uniformity effect. Secondly, the blue dots are scattered away from the centre, suggesting that the algorithm tends to overestimate the heterogeneity among products, although to a lesser extent than the first bias. Thirdly, some products are located above the red curve, especially those on the periphery, while others lie on or below it. This suggests that the quality of peripheral products tends to be overestimated.

The biases highlighted in Figure 4 are not unique to that particular session. Figure 5 presents a more general analysis that considers all 100 sessions. The left panel shows the difference between the estimated and true ratios $\frac{\hat{t}_{i2}}{\hat{t}_{i1}} - \frac{t_{i2}}{t_{i1}}$ for the I consumers, ranked from the x - to the y -axis. The estimated ratios are less dispersed than the true ratios with positive differences for consumers whose true ratio is small and negative differences for those with a high true ratio. This reflects the uniformity effect. The central panel displays the same analysis for the J products, plotting the differences $\frac{\hat{v}_{j2}}{\hat{v}_{j1}} - \frac{v_{j2}}{v_{j1}}$. Here, the estimated product types are more dispersed than the true types, although the bias is smaller than that observed for the consumer ratios. Finally, the right panel shows the distribution of $\sum_{h=1}^2 \hat{v}_{jh}^2 - 1$ for the J products, representing the algorithm's tendency to overestimate products' overall quality. The figure shows a bias against central products.

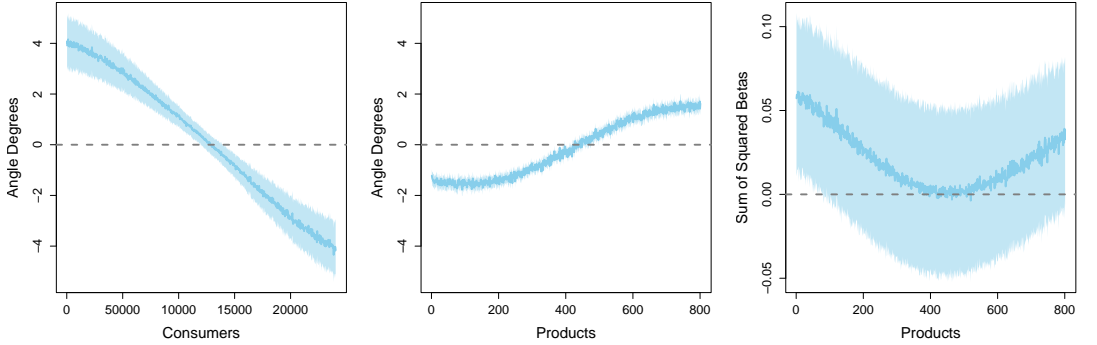


FIGURE 5.—Estimation biases with $\alpha = 0$: the consumer uniformity effect (left), the product heterogenization effect (center), and the quality bias against central products (right). The first two effects are measured by the difference between the angles, $\arctan \frac{\hat{t}_{i2}}{\hat{t}_{i1}} - \arctan \frac{t_{i2}}{t_{i1}}$ and $\arctan \frac{\hat{v}_{j2}}{\hat{v}_{j1}} - \arctan \frac{v_{j2}}{v_{j1}}$, respectively. The third effect is measured as the distance between the estimated quality $\sum_{h=1}^2 \hat{v}_{jh}^2$ and the actual quality, 1.

The quality bias that favors peripheral products could have produced a long-tail effect. However, this bias is outweighed by the uniformity effect. That is, the algorithm may think that the peripheral products are better than they actually are, but it also assumes that only a few consumers like such “extreme” products, so it recommends them infrequently. This explains the reverse long-tail effect. On the other hand, the superstar effect arises, when $\alpha = 0$, because the algorithm estimates certain products to be of higher quality than their nearby counterparts, even when all products are of the same overall quality. The products with the highest estimated qualities become the superstars that command the largest market share.

Naturally, these biases tend to disappear as the density d of the matrix of observed ratings $\tilde{\mathbf{R}}$ is large. However, in the online appendix we show that for realistic values of d , the biases remain.

8. EQUILIBRIUM PRICES

We now turn our attention to platforms such as Amazon, where consumers are charged separately for each product they purchase. We assume that sellers independently determine the prices of their products. (The platform generates revenue, for example, by collecting transaction fees.) Our objective is to examine the effect of algorithmic recommendations on equilibrium prices and consumer welfare.

Our analysis in this section is related to the extensive body of research in the field of industrial organization that has explored the impact of pre-search information on the market equilibrium. This literature remains largely unsettled. On one hand, [de Corniere \(2016\)](#) and [Zhong \(2023\)](#) demonstrate that when pre-search information directs consumers to a subset of products with the highest match values, but without ordering the products in this subset, having more information (i.e., a narrower, better-selected subset) tends to result in lower prices, unless the information becomes so precise that consumers cease searching altogether. On the other hand, [Anderson and Renault \(2000\)](#) assume that a fraction of consumers have full knowledge of their match values and demonstrate, for the case of two products, that as this fraction increases, prices rise. [Zhou \(2022\)](#), allowing for more general types of pre-search information, presents

more nuanced results. Specifically, he shows that prices decrease if the additional information does not prompt consumers to search more, but this result can be reversed if this condition does not hold.

In light of these contradictory findings, it is not clear *a priori* what the price effect of algorithmic recommendations may be. Furthermore, there are two important differences between our setup and those considered in the existing literature. First, the pre-search information in our framework consists of top-product signals that create personalized prominence, a case that, to the best of our knowledge, has not been considered thus far.³¹ Furthermore, in the papers mentioned above, preferences are idiosyncratic, whereas in our setting, as explained earlier, we have included systematic differences and similarities among consumers and products.

It turns out that algorithmic recommendations generally lead firms to increase prices, even if consumers search less. After presenting our results, we briefly discuss how they relate to previous findings in the literature.

8.1. Price equilibrium

We assume that each of the $m_{\mathcal{M}}$ products traded in market \mathcal{M} is supplied by a separate firm and that marginal production costs are zero. Firms compete in prices and correctly anticipate the demand functions.³² It is important to remember that recommendations are personalized but we assume that sellers do not engage in price discrimination.

Given the large number of products and consumers, and given that we cannot exploit symmetry, we calculate the Bertrand-Nash equilibrium prices \mathbf{p}^* numerically. The calculation is accomplished by iteratively solving the system of first-order conditions corresponding to the maximization of each firm's profit with respect to its own price.³³ The iteration starts with random prices \mathbf{p}_0^* . At each successive step $\tau = 1, 2, \dots$, starting from the vector of candidate equilibrium prices $\mathbf{p}_{\tau-1}^*$, we calculate the cut-offs $\hat{s}_{i,\mathcal{M},\tau}$,³⁴ and then from these the system of individual demand functions. For each firm j , we then calculate the new $p_{j,\tau}^*$ by maximizing firm j 's profit while holding the other prices in $\mathbf{p}_{\tau-1}^*$ constant.³⁵ In this way, we obtain the new candidate vector \mathbf{p}_τ^* . The procedure is iterated to convergence.

³¹The case where the platform recommends the product with the best estimated match value could be regarded as a limiting case of the information transmission mechanism analyzed in [de Corniere \(2016\)](#) and [Zhong \(2023\)](#), where the subset of products presented to each consumer contains just one product. In fact, however, both of those papers focus on the case where the subset contains infinitely many products.

³²The derivation of the demand functions is presented in greater detail in the online appendix. It is important to note that each firm is supposed to know perfectly which product is recommended to each consumer. This is a natural starting point for the analysis, but in future work, it would be interesting to consider the case where firms have access only to the probability that each product may be recommended to each consumer, and calculate the demand function accordingly.

³³To solve the system of first-order conditions, we use subroutine DNQSOL in the Caltech's MATH77 Fortran library, exercising special care in checking the second-order conditions.

³⁴Calculating the individually optimal cut-offs $\hat{s}_{i,\mathcal{M}}$ for our 24,030 consumers at each step of the iteration is a computationally intensive task, accounting for over 95% of the simulation time. To expedite the process, we assume that equilibrium prices follow a normal distribution with the same mean and variance as the actual equilibrium prices. We verified the validity of this assumption by also using the true distribution of equilibrium prices in some cases, but found the differences to be negligible.

³⁵This is consistent with the assumption that consumer search decisions depend on their expectations of prices and not on the actual prices, an assumption that is sometimes called *passive beliefs*.

8.2. *The impact of RSs on equilibrium prices*

Table VI presents the equilibrium prices, which are equal to the equilibrium price-cost margins under our cost normalization. The reported values are obtained by taking the average across products, with each product’s weight equal to its equilibrium market share.

α	0	0.25	0.5	0.75	1
<i>Unassisted search</i>	0.0383 (0.0000)	0.0356 (0.0000)	0.0322 (0.0000)	0.0283 (0.0000)	0.0227 (0.0000)
<i>Recommender System</i>	0.0441 (0.0002)	0.0402 (0.0001)	0.0351 (0.0000)	0.0303 (0.0000)	0.0243 (0.0000)
$\frac{RS-U_n}{U_n} \times 100$	15.10% (0.41%)	13.01% (0.28%)	9.06% (0.15%)	7.39% (0.08%)	7.13% (0.03%)

TABLE VI: Average equilibrium prices.

Algorithmic recommendations lead to an increase in prices across all values of α . The magnitude of the increase varies with α , with the largest increase (around 15%) occurring when product differentiation is purely horizontal, and the smallest (around 7%) when it is purely vertical. The online appendix confirms that varying the number of products and the level of the search cost does not modify the result. The outcome remains valid even when the algorithm is trained on data that it has previously contributed to generating.

Figure 6 provides a more detailed view of the changes in prices by showing the price changes for each individual product. Under horizontal differentiation, the price increase is relatively uniform across all products. In contrast, under pure vertical differentiation, the price of the central product slightly decreases, while the prices of other products increase. However, the decrease in the price of the central product, which carries the highest price, is offset by an increase in its market share, resulting in an increase in the weighted average price.

8.3. *Consumer surplus*

The increase in prices obviously has a negative impact on consumer welfare. However, RSs also have positive effects. First, they improve the matching between consumers and products. Second, with RSs, consumers can afford to search less extensively than in the benchmark, as the recommended product is more likely to surpass the cut-off level of surplus than a randomly chosen one. Can these positive effects ultimately outweigh the negative impact of the price rise?

To answer this question, Table VII reports the change in consumer surplus, taking into account all of these effects. Overall, consumer surplus increases. Naturally, it does not increase as much as in Table II, due to the higher prices.

The three effects on consumer surplus resulting from the introduction of RSs, namely, higher prices, reduced search, and better matching, are of the same order of magnitude. To illustrate, let us consider for instance the case where $\alpha = 0.5$. In this case, the better matching between consumers and products alone increases consumer surplus by around 1.5%. The price increase has almost the same impact but of the opposite sign, which would leave the consumer surplus almost unchanged (-0.09%) if these two effects were taken together. Thus, the overall increase of 1.3% in consumer surplus nearly coincides with the reduction in total search costs.

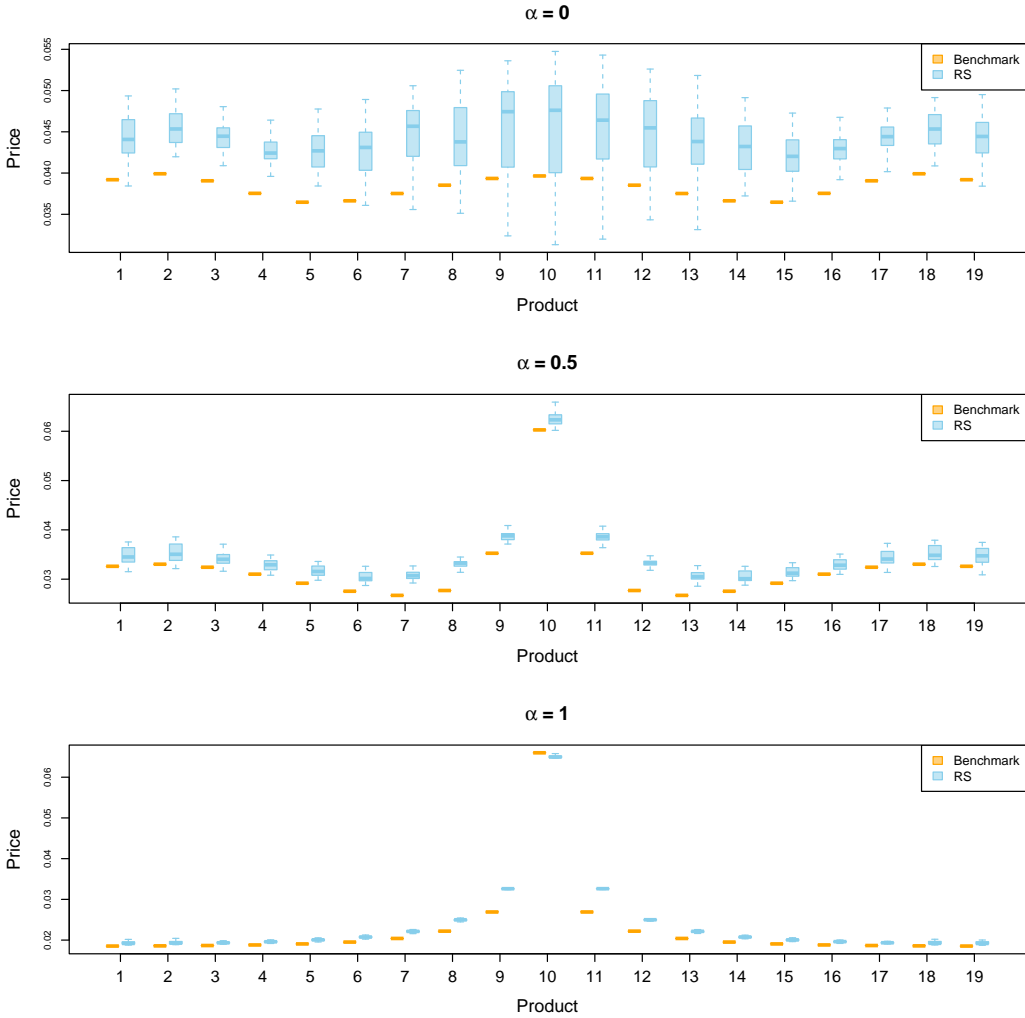


FIGURE 6.—The prices of the different products, ranked according to their location in the space of products: products 1 and 19 are peripheral, product 10 is central. The figure shows the median, inter-quartile range (IQR), and the highest and lowest value excluding outliers (obtained by subtracting 1.5 times the IQR range from the first quartile (Q1), and adding 1.5 times the IQR to the third quartile (Q3)).

The online appendix shows that the increase in consumer surplus is lower when the search cost is higher and the number of products is lower. RSs can actually decrease consumer surplus when three conditions cumulatively hold: the search cost is high, the number of products is small, and product differentiation is mainly horizontal. Taken together, however, these conditions seem rather implausible; for example, in this scenario, the fraction of consumers who follow the platform's recommendation is more than 95%. For more reasonable parameter values, the overall impact of RSs on consumer surplus is positive.

α	0	0.25	0.5	0.75	1
<i>Unassisted search</i>	0.9313 (0.0000)	0.8855 (0.0000)	0.8457 (0.0000)	0.8130 (0.0000)	0.7886 (0.0000)
<i>Recommender System</i>	0.9371 (0.0001)	0.8930 (0.0001)	0.8581 (0.0000)	0.8329 (0.0000)	0.8202 (0.0000)
$\frac{RS-U_n}{U_n} \times 100$	0.62% (0.01%)	0.85% (0.01%)	1.47% (0.00%)	2.45% (0.00%)	4.00% (0.00%)

TABLE VII: Consumer surplus, net of search costs.

8.4. Profits

Since the market is always fully covered, the increase in prices generally results in greater profits for firms. However, the impact of RSs on firms' profits is not uniform and can lead to both winners and losers, as discussed in detail in the online appendix.

8.5. Demand shifts

What drives the price changes that we have found? To gain insight, we now examine the shifts in demand generated by algorithmic recommendations. Figure 7 illustrates the overall impact of the RS on demand for both central and peripheral products across various combinations of horizontal and vertical differentiation. It appears that the way demand shifts is intricate and varies between mass and niche products, as well as with the degree of vertical differentiation, making it challenging to discern a clear pattern.

Upon closer examination, however, one can identify some major changes in the demand curves. To begin, let us focus on the first panel in the top row of the figure, which illustrates the demand for the central product under purely horizontal differentiation. We note, firstly, a *counterclockwise rotation* in the upper segment of the demand curve, specifically for prices higher than the equilibrium ones. This rotation occurs as a result of a more homogeneous customer base, which is due to the personalized nature of prominence. (If prominence were non-personalized, the rotation would be clockwise, reflecting a more diverse customer pool due to one product being recommended to all consumers.)

Secondly, for lower prices, we see an upward *vertical shift* in the demand curve. This shift occurs because RSs guide consumers toward their preferred products, increasing their willingness to pay compared to when they search individually.

Thirdly, for these lower prices, we observe a *clockwise rotation* of the demand curve. This rotation reflects the difference in demand elasticity between consumers who randomly encounter a product and those directed to their preferred products by the algorithm. When consumers can rely on algorithmic recommendations, random visits are less frequent, leading to a less elastic demand curve. Note that the last two effects dominate precisely within the price range where the equilibrium price lies.

The shifts described above are also detectable in the other panels of the figure, but they are somewhat masked by significant horizontal shifts. These shifts are positive for the central product when $\alpha > 0$, and negative for the peripheral ones. They correspond to the changes in volumes observed at zero prices and, like them, reflect the uniformity effect discussed in Section 7.4.

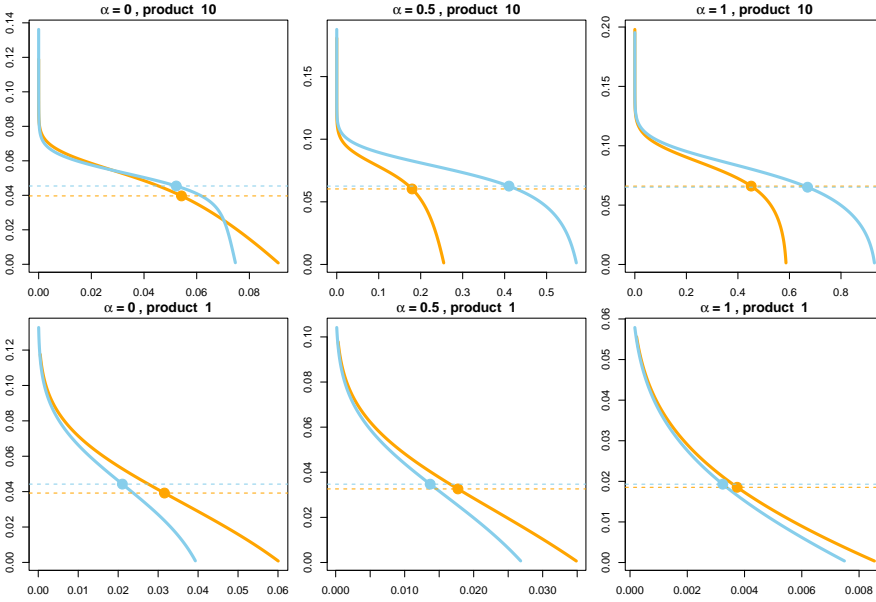


FIGURE 7.—The demand functions for the central product (product 10) and a peripheral product (product 1), for different combinations of horizontal and vertical differentiation in the benchmark (orange) and with the RS (light blue). The individual demands are calculated holding the prices of all other products constant at the equilibrium level in the benchmark case. The dots represent the equilibrium points.

Some of the shifts described above lead to price decreases, while others result in price increases. Overall, it appears that the predominant effect is the clockwise rotation in the vicinity of the equilibrium price. This corresponds to the reduction in demand elasticity highlighted by [Anderson and Renault \(2000\)](#). In other words, with personalized recommendations, consumers are more likely to find a suitable match during their initial visit, reducing their incentive to search further. This decrease in search activity diminishes competition among firms, prompting them to increase prices. However, in the case of vertically differentiated products, the increase in volume for the central product leads to a counteracting decrease in elasticity, which may result in a lower equilibrium price. In this case, the rise in the average price is primarily driven by a composition effect.

9. INFORMATION AND WELFARE

We now perform a comparative statics analysis by varying the quantity and quality of information available to the algorithms. For example, as the density of the observed rating matrix (d) increases, the algorithms access more information. Similarly, as the reporting noise (σ_e^2) decreases—or when reports are made on a finer Likert scale—the information becomes more reliable.

9.1. Varying information levels

To begin, we outline the specific modifications made to the baseline scenario to vary the information available to the algorithm.

Density. First, we allowed the density d of the observed ratings matrix $\tilde{\mathbf{R}}$, which in the baseline scenario is set to 1.2%, to vary from 0.6% to 2.4%, increasing in steps of 0.3%.

Reporting noise. Next, we varied the standard deviation of the idiosyncratic shocks ϵ_{ij} , letting it range from 0 to 40% of the standard deviation of the expected utility across the J products, with a stepsize of 2.5%. (In the baseline, it is set to 20%.)

Likert scale. A different way to vary the information derived from the reported ratings is to assume that consumers provide a value on a Likert scale, instead of a utility level. For instance, consumers might rate the product they purchased by assigning it a certain number of stars, or the platform may infer implicit, coarse ratings from the observation of consumers' behavior. To explore this possibility, we partitioned the original range of the ratings \tilde{u}_{ij} into k intervals of the same size, effectively creating a Likert scale with k levels. We assumed that the algorithm observes only the interval to which the rating belongs, instead of the exact value \tilde{u}_{ij} . We considered values of k ranging from 2 to 10.

Number of consumers and products. Perhaps less obviously, changing the number of products and consumers also affects the available information, as it alters the ratio ℓ between the number of observations and parameters to be estimated. Specifically, we explored two changes to the baseline. First, we doubled and halved both I and J while keeping their ratio constant and equal to 30. Second, we reduced the ratio of I/J to 15, and then down to 3,³⁶ while changing the levels to ensure that ℓ remains constant.

Endogenous data. As previously discussed, the quality of information deteriorates when transitioning from randomly generated data to endogenous data. Our final extension involved examining data endogenously generated as detailed in Subsection 4.2.

9.2. Measuring information levels

Detailed results for each of these extensions are provided in the online appendix and supplementary material. As it turns out, the effects of all of these changes depend more on the level of information than on the specific factors causing the change. This enables us to pool all extensions together to better elucidate the resulting pattern. The pooling capitalizes on the fact that the quantity and quality of information available to the algorithms translates into the precision of the estimates they produce and, consequently, the recommendations they make. Therefore, we consider all extensions as varying the precision of algorithmic recommendations, which we proxy by the average net utility that consumers would receive at zero prices.

9.3. Results

Let us first examine how information affects equilibrium prices. In Figure 8, the average price is plotted against the level of information, measured as just discussed. The figure indicates that having more and better information leads to higher prices, confirming that the specific way in which the change in information is generated is less important than the overall level of information available. The same general pattern holds for profits, which generally increase with the level of information.

³⁶This is motivated by the fact that the alternative dataset mentioned in footnote 22 typically have lower I/J ratios than the Netflix dataset.

Turning to consumer welfare, more and better information has two opposing effects. On the one hand, it improves the matching between consumers and products and reduces total search costs. On the other hand, we have just seen that equilibrium prices increase. When product differentiation is purely horizontal, the initial effect of improved matching prevails but is eventually dominated by the effect of higher prices, resulting in an inverted U-shaped relationship between welfare and information, as illustrated in Figure 9. (Incidentally, it appears that our baseline scenario lies on the downward slope of the curve.)³⁷

Firms have a clear interest in obtaining more information, as their profits increase monotonically with it. However, the inverted-U relationship between information and consumer surplus suggests that imposing limits on the amount of personal information that platforms can access may improve economic welfare. Therefore, some degree of privacy could be beneficial not only for its own sake but also for its positive effect on the intensity of product market competition.

The trade-off we have just identified — when consumers have access to more information, the quality of the match between consumers and products improves, but firms also gain increased market power and set higher prices — also arises in different frameworks, such as those presented in [Armstrong and Zhou \(2022\)](#) and [Jullien and Pavan \(2019\)](#). However, these papers consider frameworks without consumer search, and the mechanisms creating the trade-off seem unrelated ours. [Zhou \(2022\)](#) also presents examples where a trade-off similar to ours can arise, but, as discussed above, this occurs only if pre-search information prompts consumers to search more, not less.

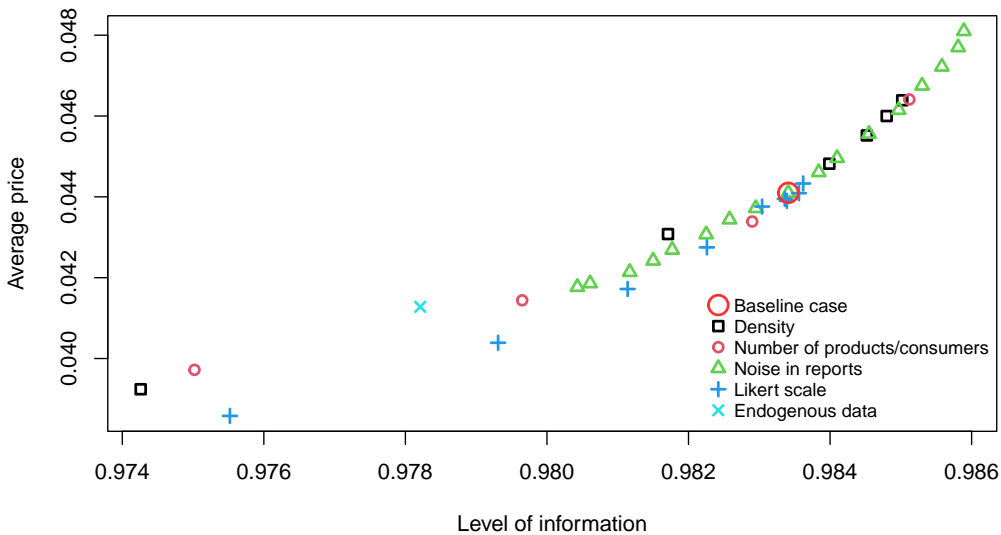


FIGURE 8.—The effect of information on equilibrium prices ($\alpha = 0$).

³⁷In the supplementary material, we show that the inverted-U pattern persists even when α is positive but not excessively large, indicating a predominantly horizontal product differentiation. In contrast, the pattern disappears when vertical differentiation predominates.

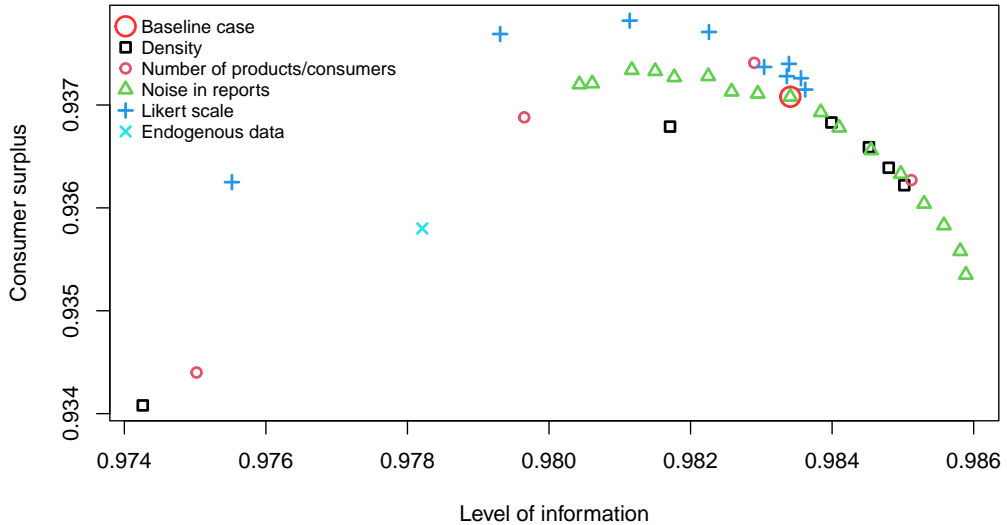


FIGURE 9.—The effect of information on consumer surplus ($\alpha = 0$).

10. SELF PREFERENCING

We now consider the scenario where the platform suggests a product other than the one with the highest estimated match value. This situation may occur when the platform promotes specific products that yield higher earnings, possibly because they are supplied by a subsidiary of the platform rather than an independent supplier.

The possibility of such self-preferencing behavior has sparked various antitrust litigations and a growing body of economic literature,³⁸ which remains largely unsettled. The framework developed in this paper offers a fresh perspective on this ongoing debate.

The specific role of algorithmic estimates when the platform engages in self-preferencing is twofold: first, for genuine recommendations when the platform does not manipulate, and second, to determine when self-preferencing should occur. Specifically, we assume that, for a given rate of self-preferencing, the platform promotes its favored product to consumers whose estimated ideal product is closest to it. This strategy minimizes potential costs incurred by the platform when manipulating recommendations.³⁹

We treat both the rate of self-preferencing, which represents the frequency with which the platform recommends its favored product to consumers for whom it is not already the best estimated match, and the platform's favored product as parameters in our analysis. Regarding

³⁸Following the seminal contribution of Hagiu and Jullien (2011), more recent papers include Calvano and Jullien (2018), de Cornière and Taylor (2019), Teh and Wright (2022), Bourreau and Gaudin (2022), Peitz and Sobolev (2022), and Bar-Isaac and Shelegia (2023).

³⁹When recommendations are misleading, consumers will revert to individual search more frequently. In such cases, consumers may end up purchasing from an alternative marketing channel, potentially resulting in revenue loss for the platform. Additionally, the platform risks losing users if their surplus falls below a certain level.

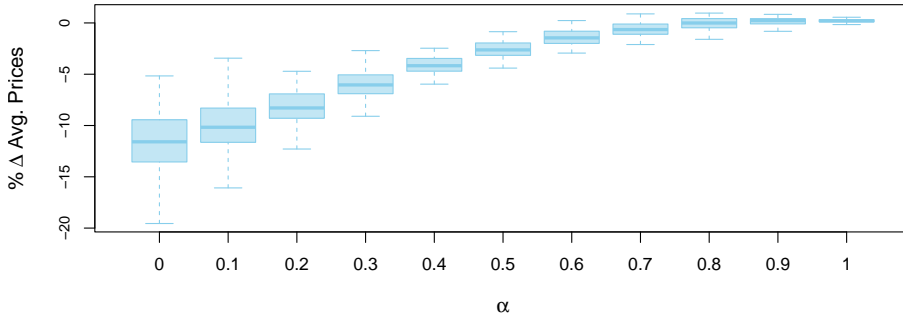


FIGURE 10.—The distribution across sessions of the impact of self-preferencing on the average price. The platform’s favorite product is the central one, and the rate of self-preferencing is 50%.

the latter, we examine two scenarios: one where the platform’s favored product is the central one (product 10),⁴⁰ and when it is one of the peripheral products (say product 1).

10.1. Intensity of competition

In all cases, self-preferencing intensifies competition and reduces equilibrium prices.

Figure 10 presents the change in equilibrium prices, relative to the case of genuine recommendations, for the case where the favored product is the central one and the rate of self-preferencing is 50%. The online appendix shows that the same qualitative pattern holds more generally. It also shows that the largest decrease in price occurs for the favored product and its closest competitors. This can be explained by the fact that without self-preferencing, the pool of consumers inspecting the favored product is relatively homogeneous. When instead the platform manipulates the recommendations, this pool becomes both larger and more heterogeneous. As a result, the demand for the favored product increases, but also becomes more elastic, leading the supplier to lower the price.⁴¹

Conversely, the demand for competing products decreases as fewer consumers are directed toward them. As a result, competitors decrease their prices to remain competitive. These firms therefore suffer both from lower volumes and lower prices, so their profits decrease. These effects are largest for the products that are closest to the favored one, as the consumers who are misdirected are those whose ideal product is relatively similar to it. For more distant products, the effect is smaller.

10.2. Profitability

Now suppose that the platform adjusts the rate of self-preferencing. Neglecting the potential costs associated with self-preferencing mentioned in footnote 39, we examine the impact on the profits of the favored products. To investigate this, we vary the rate of self-preferencing from 0% to 100% and calculate the corresponding profits. The results are depicted in Figure

⁴⁰Note that when α is large, most consumers are already directed towards the central product, regardless of self-preferencing. As a result, in this case the impact of self-preferencing is almost imperceptible.

⁴¹This is similar to the effect of uniform prominence in [Armstrong et al. \(2009\)](#).

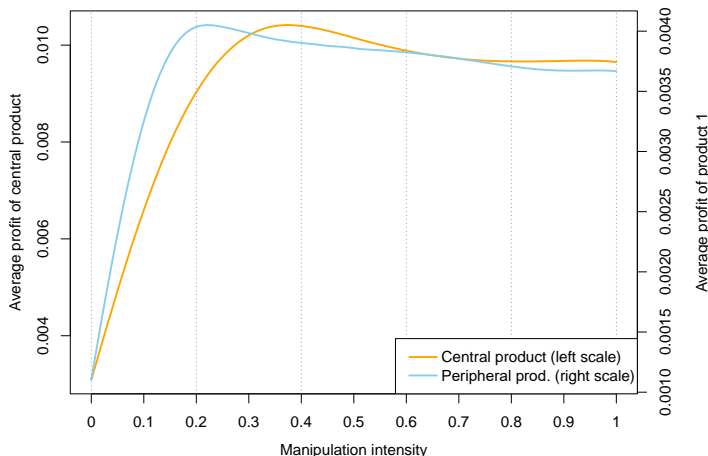


FIGURE 11.—The profits of the favored products (central in light blue, peripheral in orange) as a function of the rate of self-preferencing, with $\alpha = 0$.

11, which indicates that the profit of the favored product reaches its peak at relatively low rates of self-preferencing, roughly between 20% to 30%.

From the figure, it becomes apparent that there is a trade-off between price and volume: increasing self-preferencing leads to higher demand for the favored product, but it also intensifies competition, resulting in lower prices. The observed decline in profit as the rate of self-preferencing increases is driven precisely by the competition-enhancing effect highlighted above.

10.3. Consumer surplus

Consumer surplus decreases when recommendations are manipulated relative to the case where they are sincere. Due to the price reduction, however, the decrease is relatively small. The analysis presented in the online appendix suggests that, given a certain market structure, the impact of self-preferencing on consumer surplus is likely limited. However, a more concerning issue could be the decrease in competitors' profitability, which may lead to exits or deter entry in the long run. As a result, policymakers may view self-preferencing as more of an exclusionary abuse rather than an exploitative one.

11. CONCLUSIONS

We have proposed a novel framework to analyze the potential effects of algorithmic recommendations on competition in product markets, applying this framework to three significant economic issues: the possibility that RSs could increase market concentration; enable sellers to more effectively segment the market and charge higher prices; and allow platforms to manipulate their recommendations for their own benefit.

While our analysis confirms these risks, it also demonstrates that RSs exhibit pro-competitive effects, such as improving the match between products and consumers and reducing the need for costly search. Based on our quantitative assessment, the pro-competitive

effects of RSs are at least as significant as their anti-competitive effects, and in many cases, outweigh them. Nevertheless, we also find that increasing the amount of information available to RSs may harm consumers. Therefore, imposing limits on the platforms' access to personal information may be socially desirable. In addition to safeguarding privacy, such limitations could promote competition among sellers, leading to higher consumer surplus.

Methodologically, the analytical framework proposed is flexible and can be adjusted to incorporate different algorithms, preferences, and assumptions about the search process to deliver more robust conclusions. We conclude by mentioning some potential extensions that could be explored in future work.

Firstly, we have assumed that consumers know the equilibrium price distribution but not individual product prices. While this approach is widely used in search theory, in online markets, consumers can easily obtain information on product prices. Accordingly, a more recent search literature has explored the case where consumers initially know the prices of individual products but not how well each product fits their personal preferences. Extending our analysis to this case necessitates two adjustments. First, in the individual search benchmark, we must account for the possibility of ordered search, where the sequence of products searched is determined by their prices. Second, for consistency, the platform should recommend the product with the highest estimated net surplus ($\hat{r}_{ij} - p_j$) rather than the highest estimated utility (\hat{r}_{ij}). Both modifications lead to lower prices and must be implemented simultaneously to ensure valid comparisons. While the second adjustment is straightforward, the first significantly increases the complexity of the analysis. As previously noted, calculating the cut-offs $\hat{s}_{i,\mathcal{M}}$ is already the most computationally intensive part of the numerical analysis, and adding this layer of complexity makes the process even more demanding.

Secondly, when a platform recommends several products instead of just one, and presents them in a particular order, it complicates the analysis of consumer behavior with RSs (while the individual search benchmark does not change). In this scenario, consumers may have a stronger incentive to continue searching after inspecting the first product because further sampling would not be random. This increased search effort is likely to be pro-competitive.

Thirdly, two or more platforms often compete to attract consumers. In this case, the quality of the algorithm's estimates may represent a factor of competitive advantage or disadvantage.

In more extended models, one could analyze the impact of RSs on other economic decisions, such as entry and exit, product quality, and R&D investment. Although these other aspects are beyond the scope of this study, our analytical framework can be used to systematically analyze them in future research.

In terms of policy, our framework can be used to assess the effect of various forms of regulations. For example, it may be interesting to consider a policy that limits the level of personalization in recommendations. Although this could potentially decrease the match-value of the recommendation system, it could also help mitigate the price increases we have found.

Finally, an intriguing and challenging extension to our analysis would be to consider the possibility that RSs may influence individual preferences. This possibility arises because tastes are endogenous, a fact that has been recognized at least since [Knight \(1923\)](#) but has been rarely studied analytically. The act of inspecting a product, such as starting to watch a movie, may not only reveal the match value but also shape consumer preferences. If this is the case, it raises concerns about the potential for platforms to manipulate individual tastes to their advantage.

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ONLINE APPENDIX

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The appendix offers additional results and detailed modeling information that were not included in the main text. To maintain coherence, we have preserved the section numbering to align with the corresponding sections in the main article. Consequently, Section A.x.y complements Section x.y in the main text and so forth.

3. MODEL-BASED RECOMMENDER SYSTEMS

3.1. *Estimation procedure: regularization and cross-validation*

In this section we discuss more details of the estimation procedure of the RS, its regularization and the cross-validation routine mentioned in Section 3.3 of the paper.

The minimization of the loss function (6) in the paper proceeds by calculating the gradient vector of partial derivatives w.r.t. the $J \times H + I \times H$ decision variables and using numerical methods to find improved values of $\hat{\mathbf{T}}, \hat{\mathbf{V}}$.

In vector notation, the gradient is

$$\mathbf{G} = - \begin{pmatrix} \mathbf{E}\hat{\mathbf{V}} - \lambda_t \hat{\mathbf{T}} \\ \mathbf{E}'\hat{\mathbf{T}} - \lambda_v \hat{\mathbf{V}} \end{pmatrix} \quad (1)$$

where \mathbf{E} is the $I \times J$ matrix of residuals $e_{ij} = \tilde{r}_{i,j} - \sum_{h=1}^H \hat{t}_{ih} \hat{v}_{jh}$.

The estimation algorithm is a version of the Stochastic Gradient Descent algorithm that is available upon request. Given some arbitrary initial values for $\hat{\mathbf{T}}_0, \hat{\mathbf{V}}_0$, at iteration t the algorithm computes an estimate of \mathbf{G}_t and iterates according to:

$$\begin{pmatrix} \mathbf{T}_{t+1} \\ \mathbf{V}_{t+1} \end{pmatrix} = \begin{pmatrix} \mathbf{T}_t \\ \mathbf{V}_t \end{pmatrix} - \eta \mathbf{G}_t, \quad (2)$$

where $\eta > 0$ is the step size and $\mathbf{T}_{t+1}, \mathbf{V}_{t+1}$ are the new values of the variables of interest at the end of the iteration. The procedure is stopped at 50 iterations, to avoid overfitting and reduce the computational burden.

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The regularization parameters λ_t , λ_v and η and the number of factors H when not pre-specified (more on this below), are determined with a 5-fold cross-validation, which is customary in machine learning.

In particular, we partition the ratings into five equally sized subsets. We then use four of these sets to estimate the model and compute the out-of-sample prediction error for the fifth set, obtaining a measure of the predictive ability of the model (RMSE). For a given parameterization, this procedure is replicated changing the left-out set, and the average RMSE is computed.

The entire procedure is then replicated for a range of possible values of λ_t , λ_v and η (as well as H when it is not pre-specified). Specifically, we have used the grid $\lambda_t, \lambda_v, \eta \in \{0.01, 0.05, 0.1, 0.15, 0.20\}$. We then choose the values of the hyper-parameters that minimize the average RMSE as computed above.

In one of the robustness checks, we allow the algorithm to choose the number of latent factors by means of a cross-validation procedure (results are reported in supplementary material). In this case, holding the true number of latent factors equal to $H = 2$, we let the algorithm choose the estimate \hat{H} from a grid 1, 2, 3, 4, 5¹. Table 1, which reports the fraction of sessions where the number of latent factors is estimated at each of the possible five values, shows that the cross-validation routine almost systematically identifies the correct number of factors, $H = 2$. In less than 5% of the cases, the algorithm chooses $\hat{H} = 3$ instead.

H/α	0	0.25	0.5	0.75	1
1	0.00	0.00	0.00	0.00	0.00
2	0.95	0.96	0.95	0.96	0.97
3	0.05	0.04	0.05	0.04	0.03
4	0.00	0.00	0.00	0.00	0.00
5	0.00	0.00	0.00	0.00	0.00

TABLE 1. Each cell reports the fraction of cases in which the cross-validation procedure identifies a number of factors H as indicated in each rows, for different values of α (in the columns).

4. PRODUCTS AND PREFERENCES

4.1. Preferences and transportation costs

As mentioned in the main text, our model of horizontal product differentiation shares certain similarities with the conventional Hotelling framework. To emphasize these resemblances, we calculate the “transportation-cost” implicit in our framework when $\alpha = 0$ and compare it to the cost associated with more traditional Hotelling frameworks.

To calculate the implicit transportation cost, we employ polar coordinates as in Perego (2020), denoting product positions as $\{t_{i1} = \cos \theta, t_{i2} = \sin \theta\}$ and consumer positions as $\{v_{j1} = \cos v, v_{j2} = \sin v\}$. Using these coordinates, a product’s location is represented by v within the range $[0, \frac{\pi}{2}]$, and a consumer’s location is represented by θ within the same range. Therefore, $\theta - v$ represents the angular distance between a consumer located at θ and a product located at v . The utility that a consumer at θ derives from a product at v is given by:

$$u(\theta, v) = \cos \theta \cos v + \sin \theta \sin v = \cos(\theta - v). \quad (3)$$

¹In another robustness check we set the number of latent factors to $H = 5$. In this case we assume the algorithm knows the true number of latent factors. Results are in section ??.

Since the utility that consumers receive from their ideal product (where $v = \theta$) is 1, the “transportation cost” is simply $1 - \cos(\theta - v)$; there is no explicit transportation cost parameter.

Figure 1 contrasts our transportation cost with a quadratic cost.

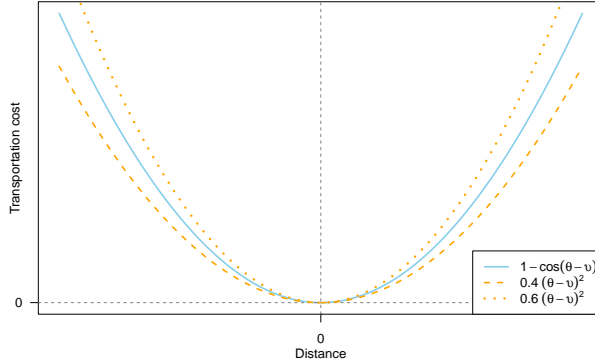


FIGURE 1.—The continuous curve depicts the transportation cost $1 - \cos(\theta - v)$ in our model. For comparison, the figure also presents the standard quadratic transportation cost $\tau(\theta - v)^2$. The dashed curve corresponds to $\tau = 0.4$, and the dotted one to $\tau = 0.6$. For $\tau = 0.5$, the quadratic transportation costs are almost indistinguishable from that of our latent factor model.

6. MARKET CONCENTRATION

6.3. *The superstar effect with horizontal differentiation*

Table IV shows that the superstar effect is significant when vertical differentiation is important (large α), but not when product differentiation is mainly horizontal (small α). In fact, RS creates superstars even under horizontal product differentiation. This tendency is not apparent in Table IV because with $m_{\mathcal{M}} = 19$ products there are many products that are “central,” and the ones selected by the RS as superstars vary randomly from session to session. To eliminate this effect, we rank products in terms of their market share and calculate the average share, across the 100 sessions, of the most popular product, the second most popular, and so on. We do so for the case with RS and for the benchmark of unassisted search. The resulting distribution is depicted in Figure 2.

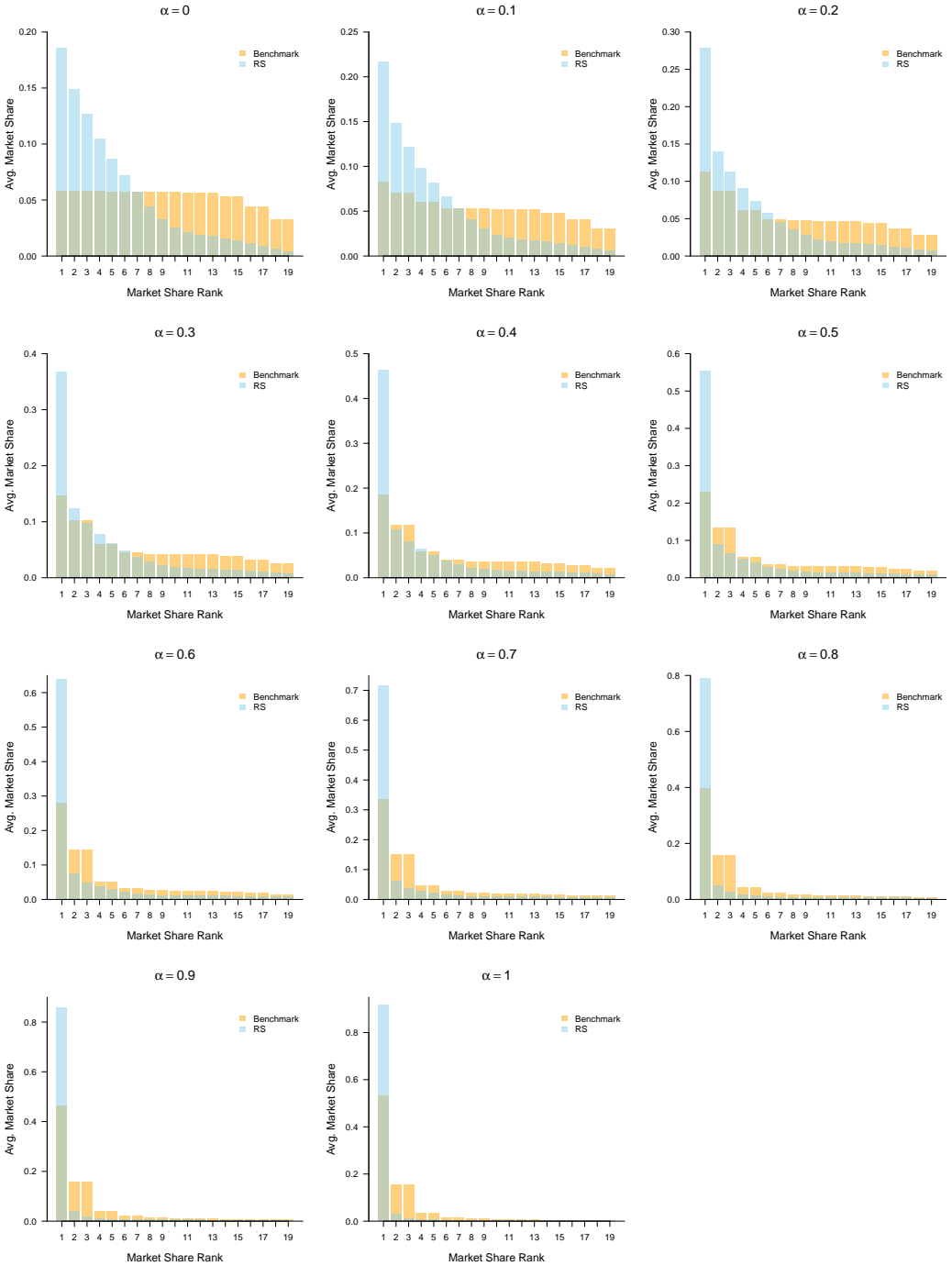


FIGURE 2.—Average Market share of products ordered by popularity, for different degrees of product differentiation ranging from pure horizontal ($\alpha = 0$), and pure vertical ($\alpha = 1$). The average is computed across session.

In the individual search benchmark, all products except for the most peripheral ones have similar market shares of around 5% when $\alpha = 0$. However, with the RS, the most popular product has an average market share of over 15%, indicating the presence of a significant superstar effect even when $\alpha = 0$. This effect becomes even stronger as α increases.

Reconsidering the definition of a ‘superstar’ as the product with the highest market share, we can identify this product in Figure 3. The figure illustrates the frequency with which each product attains superstar status across 100 simulations. Products are organized based on their position within the product space, spanning from the x - to the y -axis, as previously shown in Figure 1. Notably, central products exhibit a greater tendency to become superstars even in cases where $\alpha = 0$. As α increases, a noteworthy trend emerges: the product located at the center of the product space (characterized by $m_{\mathcal{M}} = 19$ and product number 10) is selected more frequently. In fact, for values of α equal to or exceeding 0.3, this specific product consistently assumes the role of the superstar.

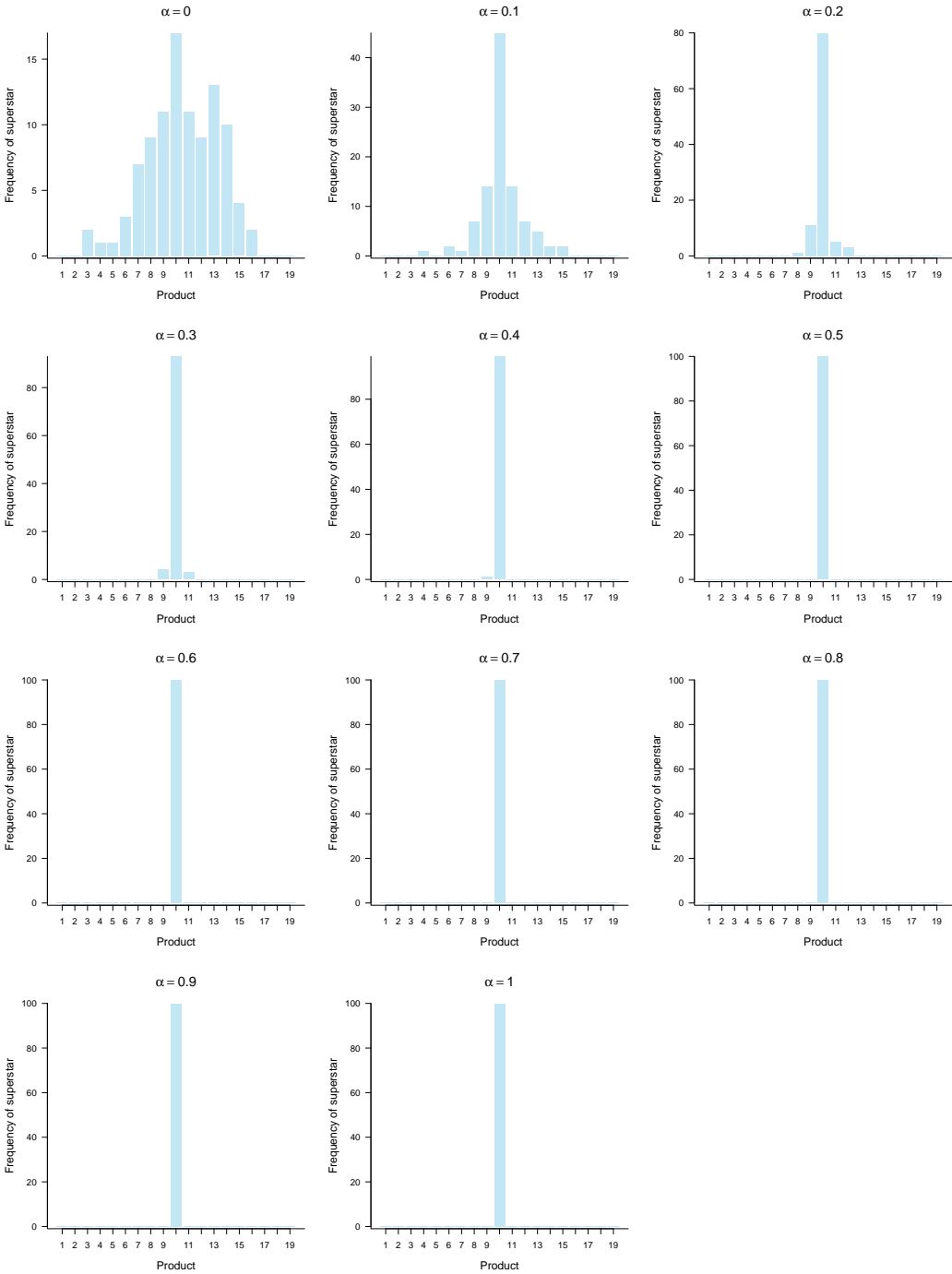
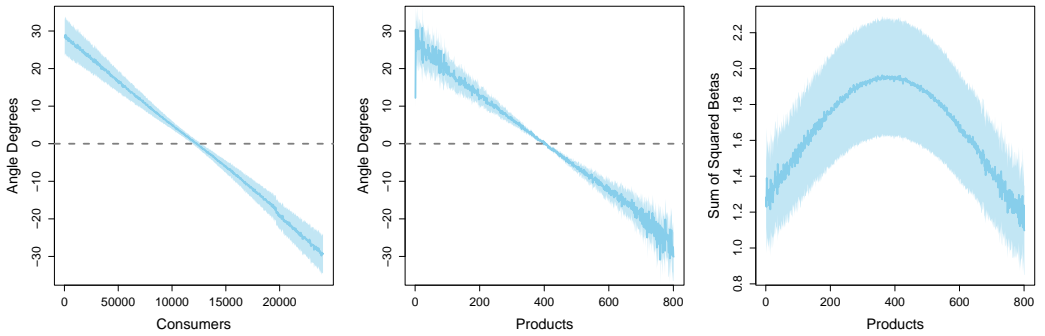


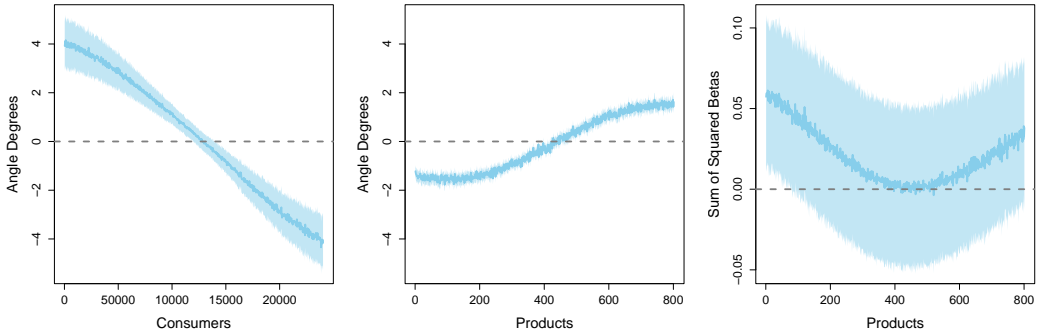
FIGURE 3.—Frequency with which each product is selected as the superstar. Products are ordered on the basis of their location in Figure 1 (body of the paper), from the x -axis (product 1) to the y -axis (product 19). The central product is product 10.

6.4. *The uniformity effect*

In Figure 4, we study the uniformity effect for different values of the density d of the rating matrix. In particular, in examining the estimation biases across varying values of d , a noteworthy pattern emerges. Except for the lowest value of d , the biases appear to be consistently similar for θ , β , and the sum of squares of β . This consistent behavior suggests that the estimation process remains stable for these parameters over a range of d values. However, at the lowest d , instead of observing increasing heterogeneity in the estimated β , as seen in all other cases, there is a homogenization of estimates of β . This homogenization aligns with what is consistently observed for estimates of θ . Furthermore, for this lowest d , there's a bias in favor of central products. In contrast, for all other d values, there's a bias leaning towards niche products.



(a) Estimation biases with density $d = 0.6\%$.



(b) Estimation biases with density $d = 1.2\%$.

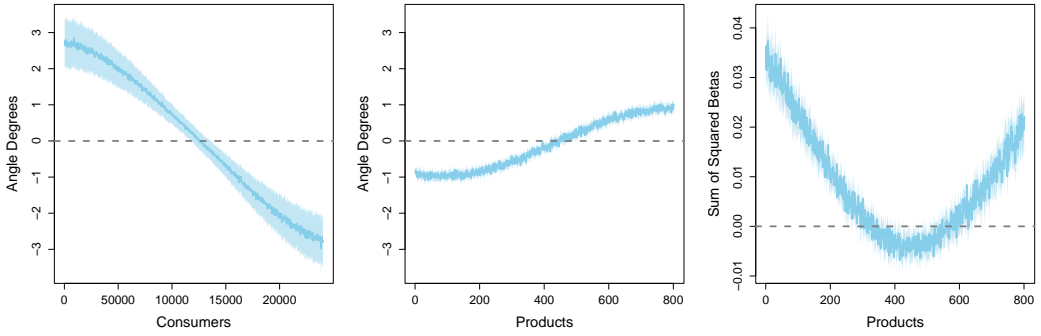
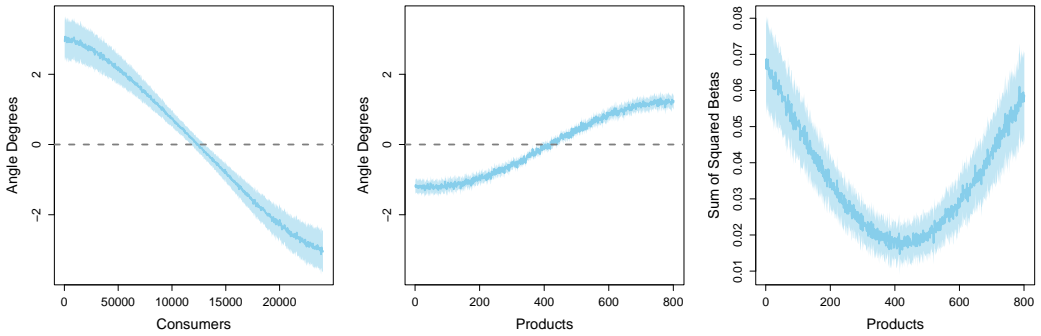
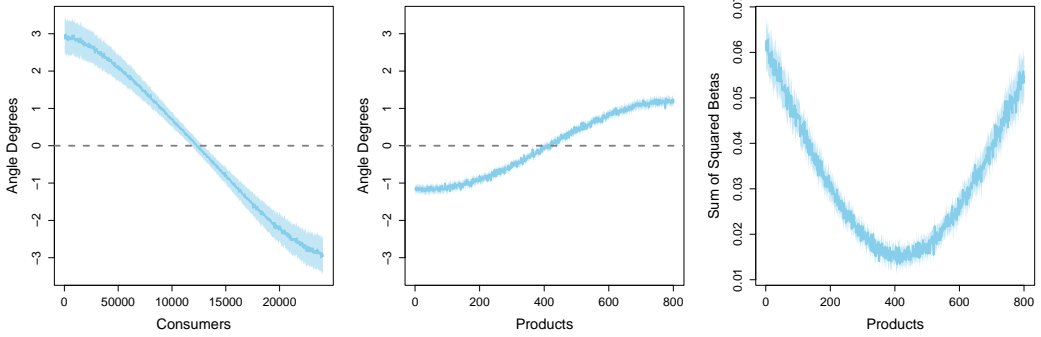
(c) Estimation biases with density $d = 2.4\%$.(d) Estimation biases with density $d = 3.6\%$.(e) Estimation biases with density $d = 4.8\%$.

FIGURE 4.—Estimation biases for various densities.

7. EQUILIBRIUM PRICES

7.1. Deriving demand functions

The demand functions are calculated as follows. Let q_{ij} denote the probability that the surplus provided by product j to consumer i surpasses the cutoff $\tilde{u}_{ij} - p_j$. The probability x_{ij} that consumer i purchases product j then is:

$$x_{ij} = \frac{1}{n_{\mathcal{M}}} q_{ij} + \left[\sum_{\ell \in \mathcal{M}} \frac{1}{n_{\mathcal{M}}} (1 - q_{i\ell}) \right] x_{ij} \quad (4)$$

The first term on the right hand side is the probability that product j is sampled first and is immediately purchased. The term inside square brackets is the probability that no product is purchased after the first search, in which case the search process starts anew. Solving this equation yields:

$$x_{ij} = \frac{q_{ij}}{\sum_{\ell \in \mathcal{M}} q_{i\ell}}. \quad (5)$$

The total demand for product j is obtained by aggregating individual demands x_{ij} across consumers. Since q_{ij} decreases with p_j , this yields a system of decreasing demand curves.

With RSs, the algorithm recommends to consumer i the product

$$j^*(i, \mathcal{M}) = \arg \max_{j \in \mathcal{M}} \hat{r}_{ij}. \quad (6)$$

The demand for this product is:

$$x_{ij^*(i, \mathcal{M})} = q_{ij^*} + (1 - q_{ij^*}) \frac{q_{ij^*}}{\sum_{\ell \in \mathcal{M}} q_{i\ell}}. \quad (7)$$

whereas the demand for products $j \neq j^*$ is

$$x_{ij} = (1 - q_{ij^*}) \frac{q_{ij}}{\sum_{\ell \in \mathcal{M}} q_{i\ell}}. \quad (8)$$

7.2. Individual profits

In section 7 of the main text, we mention the effect of the RS on equilibrium profits. Figures 5 and 6 illustrate these results in detail, showing how the RS affects the profits earned by each of the 19 firms in the market. The figure shows that when $\alpha = 0$, the peripheral firms definitely lose, while the gains and losses for the other firms do not seem to follow a clear pattern. When $\alpha > 0$, on the other hand, the central firm emerges as the clear winner, while the losers include not only the peripheral firms but also the closest competitors of the central firm. For the remaining firms, the effect is not significantly different from zero. In the figure, the median profit of central products decreases because creating champions with highly asymmetric market shares, the RS sometimes benefits and other times disadvantages those products. However, as anticipated from the price hikes, the average profits significantly increase, as show in Table 2.

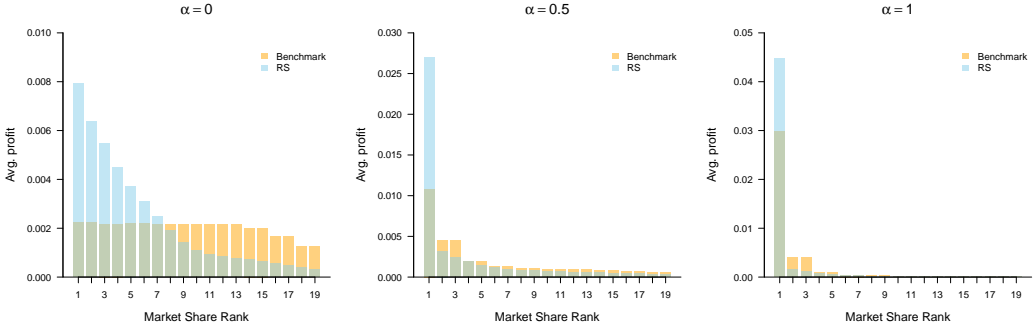


FIGURE 5.—The average (across sessions) profits of the suppliers ranked by market share.

α	0	0.25	0.5	0.75	1
<i>Unassisted Search</i>	0.0022 (0.0000)	0.0053 (0.0000)	0.0108 (0.0000)	0.0192 (0.0000)	0.0298 (0.0000)
<i>Recommender System</i>	0.0031 (0.0003)	0.0148 (0.0004)	0.027 (0.0002)	0.0372 (0.0001)	0.0448 (0.0001)
$\frac{RS - Un}{Un} \times 100$	40.91%	179.25%	150.00%	93.75%	50.34%

TABLE 2
AVERAGE PROFITS

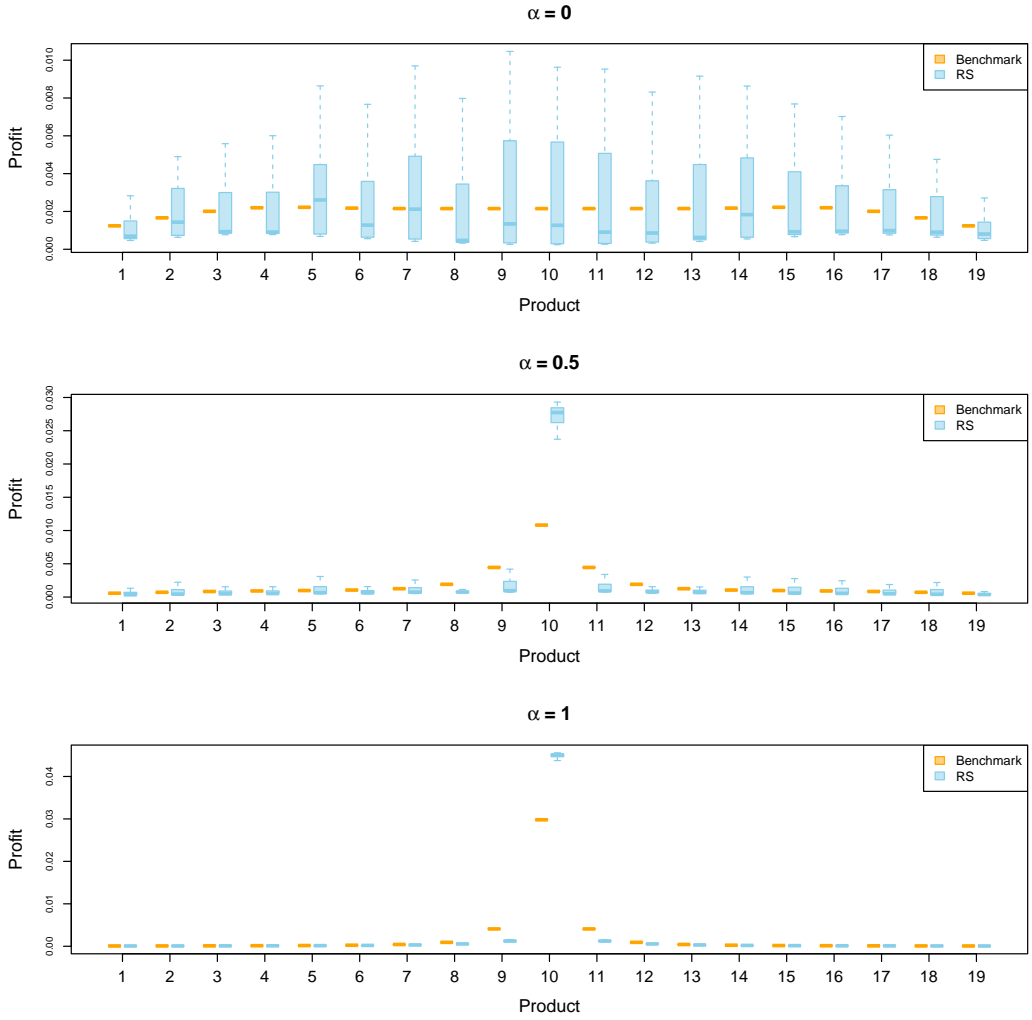


FIGURE 6.—The profits of the suppliers of the different products, ranked from 1 to 19 according to their location in the space of products: products 1 and 19 are peripheral, product 10 is central. As customary, the figure shows median, inter-quartile range (IQR), and the values obtained by subtracting 1.5 times the IQR range from the first quartile (Q1), and obtained by adding 1.5 times the IQR to the third quartile (Q3). That is the highest and lowest value excluding outliers).

7.3. Number of products and search cost

In this section, we illustrate how our results change when we vary the number of products $m_{\mathcal{M}}$ and the level of the search cost c_s . We consider both the effect on market concentration when prices are exogenously set at zero, and the effect on equilibrium prices and consumer surplus.

Specifically, we consider three possible levels of the search cost: $c_s = 0.002$, $c_s = 0.004$ (the baseline value considered in the main text), and $c_s = 0.006$. For each of these values, and for each variable of interest, we present a heat map obtained by varying the number of products $m_{\mathcal{M}}$ from 7 to 31 (the baseline value considered in the main text being $m_{\mathcal{M}} = 19$), and the degree of vertical product differentiation α from 0 to 1. The heat maps represent the percentage change in the relevant variables with RSs relative to the individual search benchmark.

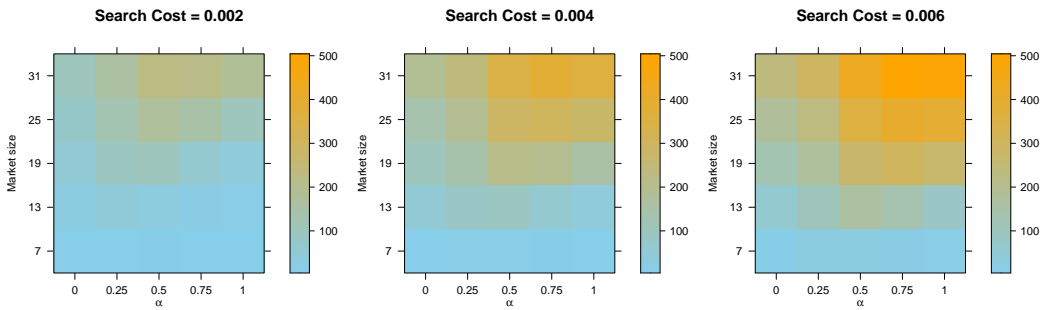


FIGURE 7.—Percentage shift of the HHI Index from Benchmark to RS.

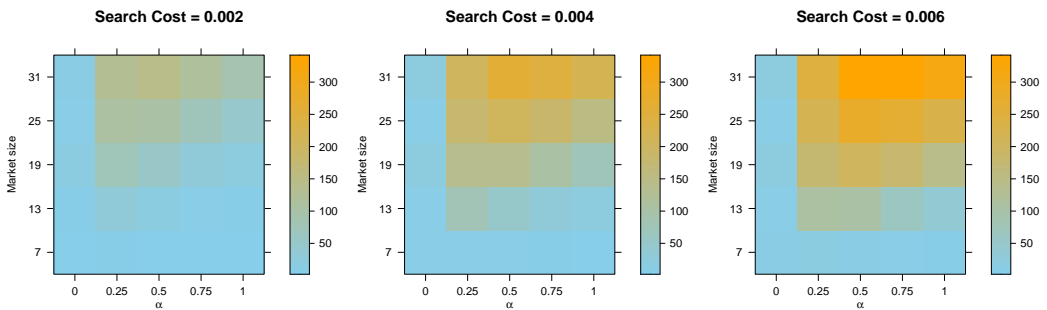


FIGURE 8.—Percentage shift of the market share of the central product from Benchmark to RS.

We have seen in the main text that the impact of the RS on market concentration, as measured by the HHI, increases with α . Figure 7 confirms this finding and shows that the effect also increases with the number of products and the level of the search cost. Unsurprisingly, the superstar effect (Figure 8) and the reverse long-tail effect also exhibit a similar pattern.

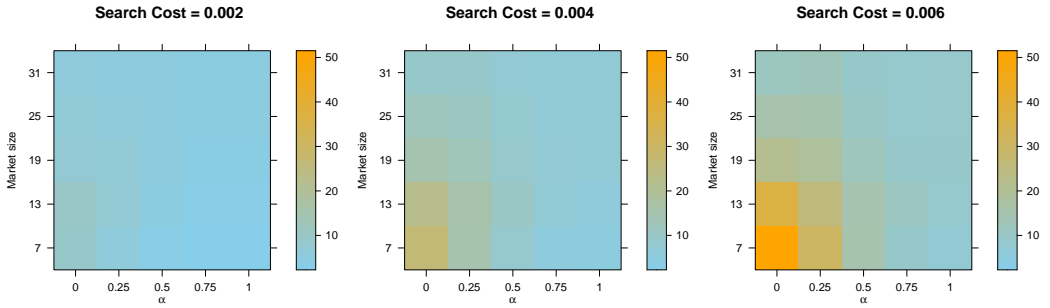


FIGURE 9.—Percentage shift of equilibrium prices (average across products) from benchmark to RS

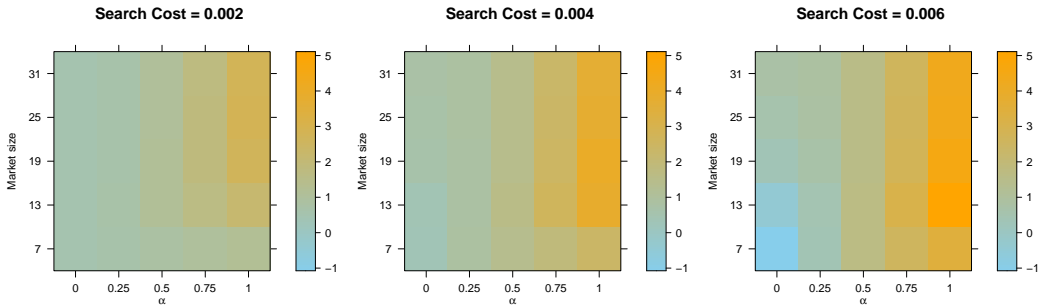


FIGURE 10.—Percentage shift of consumer surplus net of search costs from benchmark to RS.

In the main text, we have seen that the effect of the RS on equilibrium prices is highest when product differentiation is mainly horizontal (small α). Again, Figure x confirms this finding. Furthermore, it shows that the increase in equilibrium prices increases with the search cost and decreases with the number of products. When the search cost is high and the number of products is small, the average price may increase by as much as 50%.

In the paper, we observed that the RS has the most significant impact on equilibrium prices when product differentiation is primarily horizontal. Figure 9 reaffirms this. Additionally, the figure illustrates that the rise in equilibrium prices is proportional to the search cost and inversely related to the number of products. In cases where the search cost is high and the number of products is limited, the average price may experience an increase of up to 50

Figure 10 shows that this large increase in the average price may more than offset the positive effect of the RS on the matching of consumers and products, lading to a decrease in the level of consumer surplus. This possibility arises when the search cost is high ($c_s = 0.006$), product differentiation is horizontal ($\alpha = 0$), and the number of products is small ($m_{\mathcal{M}} = 7$ or 13).

9. BIASED RECOMMENDATIONS

In this section we provide further information for the case in which the RS provides manipulated recommendations, as discussed in the main text.

9.1. Intensity of competition

We will now dissect the average impact of manipulation on prices, illustrating its effects on various products while varying the favored product's identity. Figures 11-13 complement the main paper's findings by presenting results with a 50% manipulation rate, where the central product is favored, for different values of α .

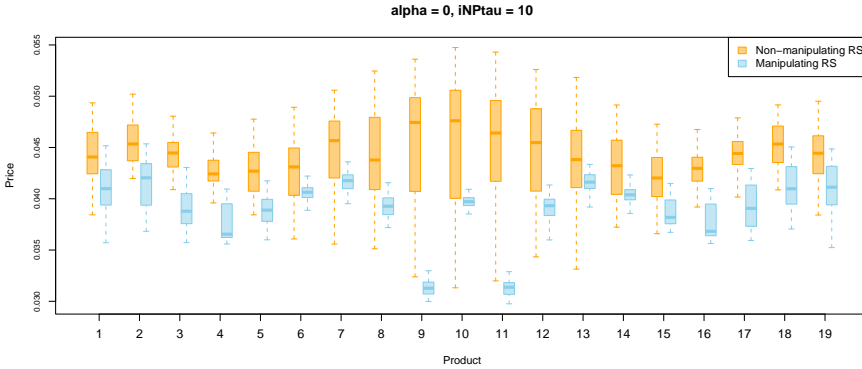


FIGURE 11.—Impact of 50% manipulation on prices for all products when product 10 is favored, with $\alpha = 0$.

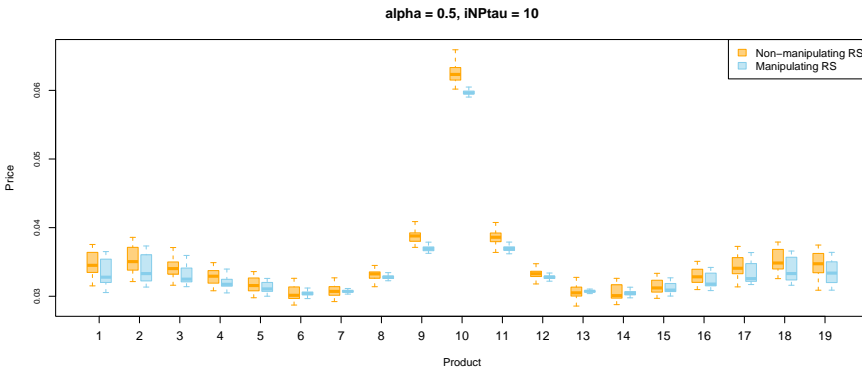


FIGURE 12.—Impact of 50% manipulation on prices for all products when product 10 is favored, with $\alpha = 0.5$.

The next figures 14, 15, 16 consider the same analysis of the effects of manipulation on prices, this time varying the favored product's identity. In all cases, manipulation intensifies competition and reduces equilibrium prices.

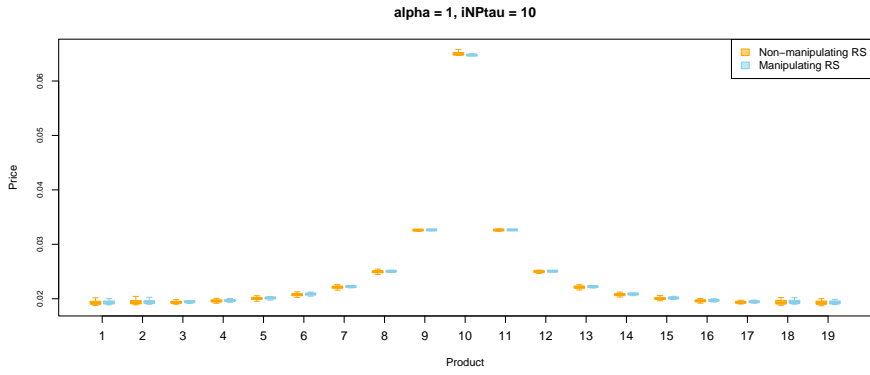


FIGURE 13.—Impact of 50% manipulation on prices for all products when product 10 is favored, with $\alpha = 1$.

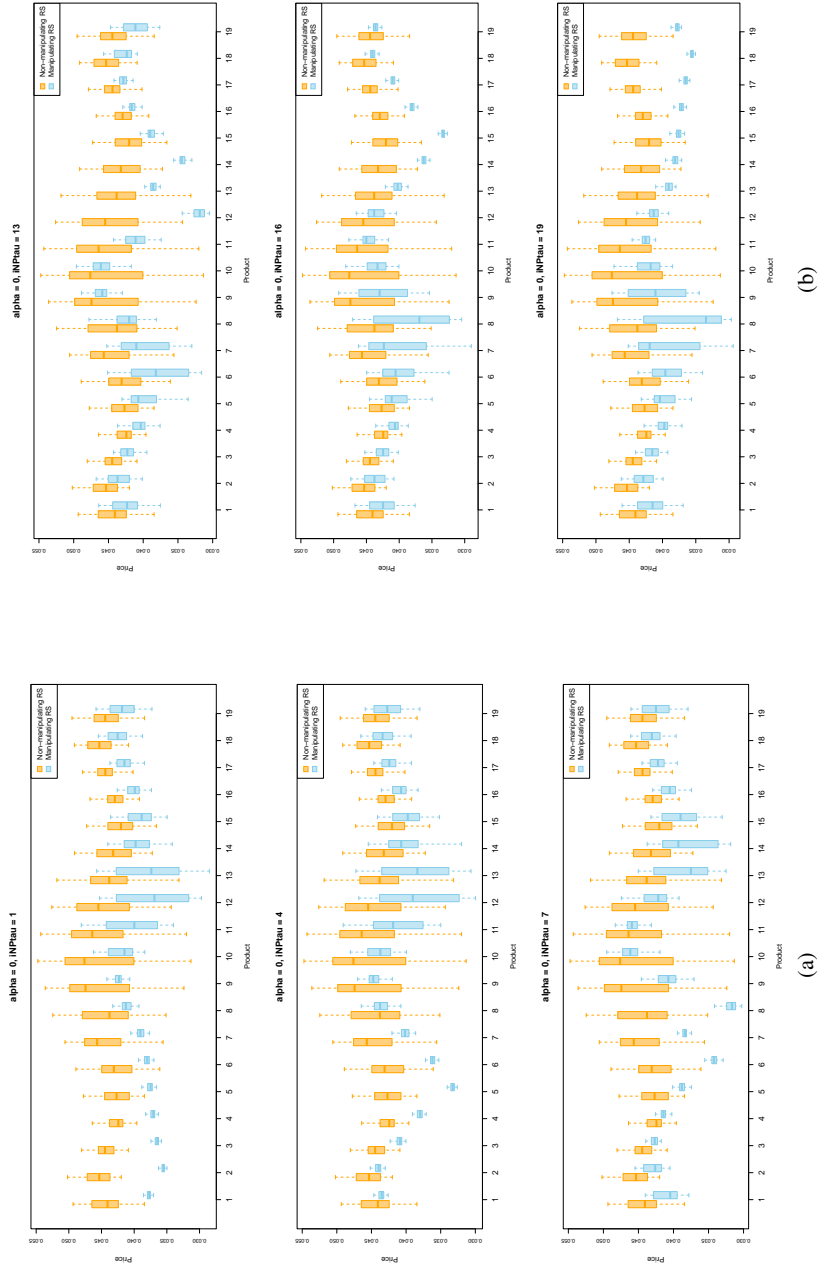


FIGURE 14.—Effect of manipulation on equilibrium prices, when the favored product differs from the central one. ‘INPtau’ is the favored product’s identity. Our labeling convention designates products based on their location, with products 1 and 19 as peripheral and product 10 as central. This figure: parameters $\alpha = 0$, and manipulation frequency maintained at 50%.

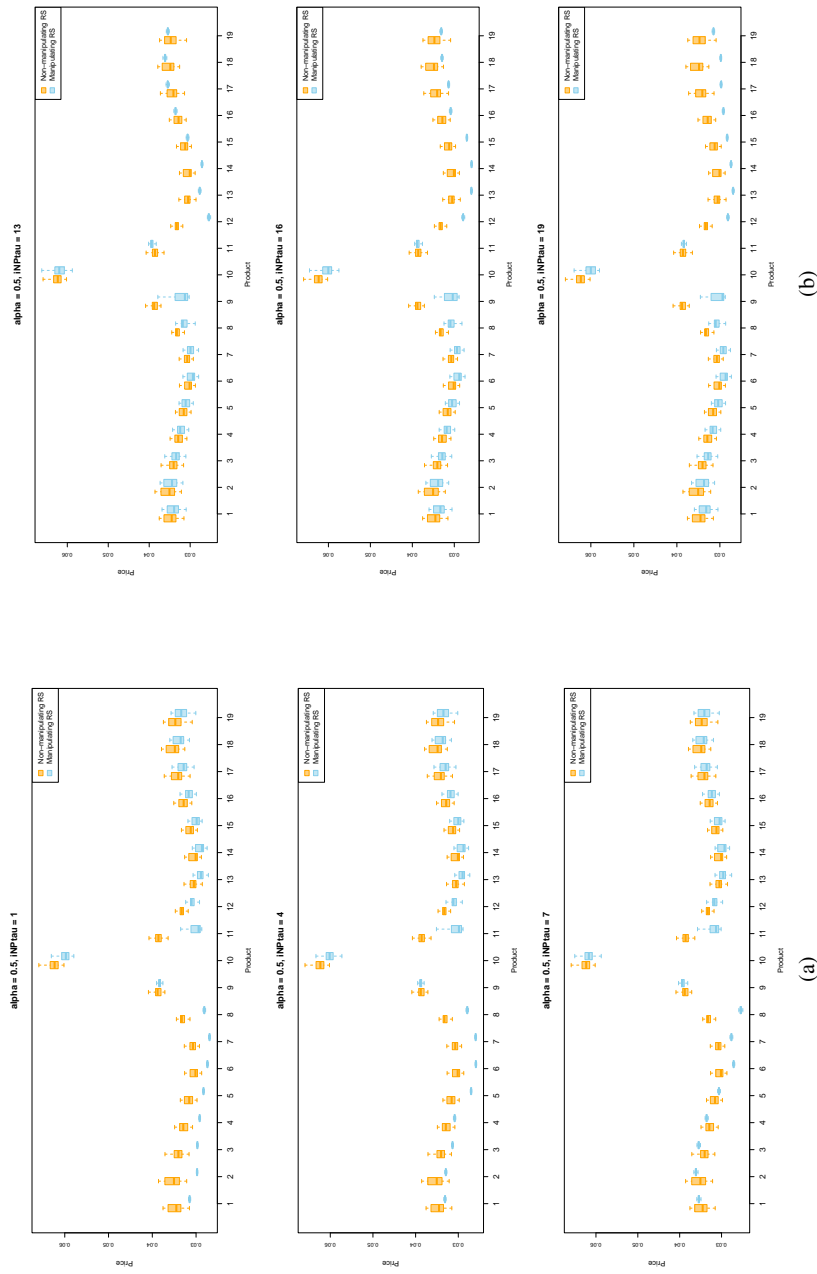


FIGURE 15.—Effect of manipulation on equilibrium prices, when the favored product differs from the central one. ‘INPtau’ is the favored product’s identity. Our labeling convention designates products based on their location, with products 1 and 19 as peripheral and product 10 as central. Key this figure: parameters $\alpha = 0.5$, and manipulation frequency maintained at 50%.

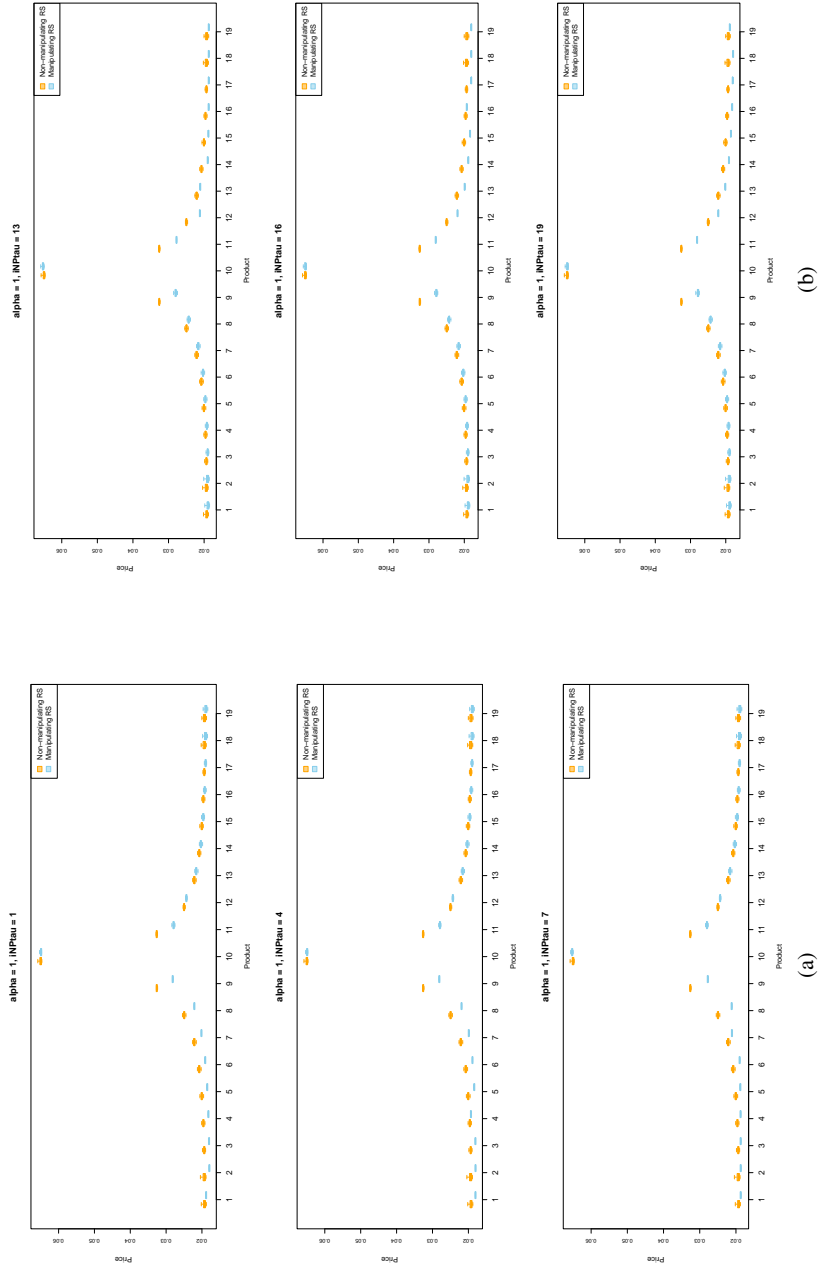


FIGURE 16.—Effect of manipulation on equilibrium prices, when the favored product differs from the central one. ‘INPtau’ is the favored product’s identity. Our labeling convention designates products based on their location, with products 1 and 19 as peripheral and product 10 as central. Key this figure: parameters $\alpha = 1$, and manipulation frequency maintained at 50%.

9.2. Profits

We here unpack the effect of manipulation on profits by showing the impact on the different products and by varying the identity of the favored product. Figure 17 complements the main paper’s findings by presenting results with a 50% manipulation rate, where the central product is favored. Manipulation has a strong positive effect on the favorite product and a negative effect on neighboring products that diminishes with the distance.

In analogy with the analysis on prices, we repeat the exercise by changing the identity of the favored product (figures 20, 21 and 22) In all instances, manipulation heightens competition and diminishes equilibrium profits.

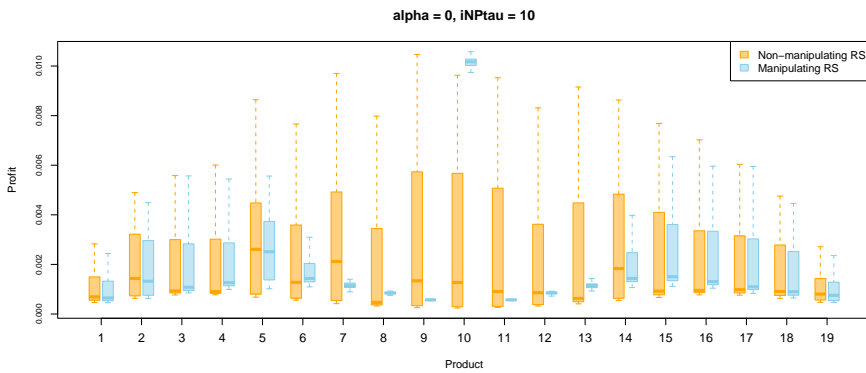


FIGURE 17.—Impact of 50% manipulation on profits for all products with product 10 favored by the recommender system ($\alpha = 0$).

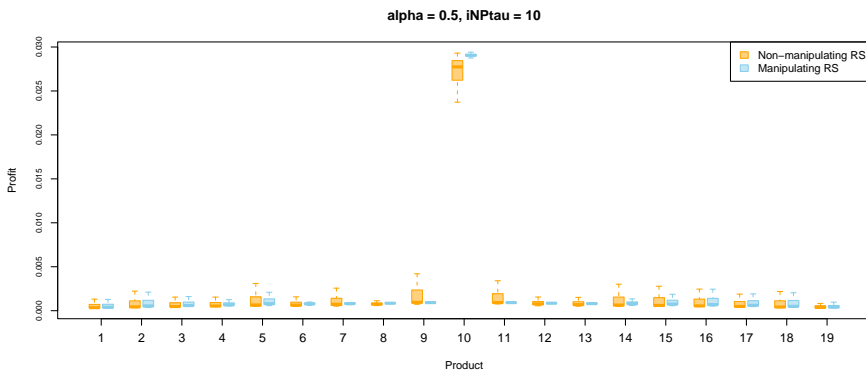


FIGURE 18.—Impact of 50% manipulation on profits for all products with product 10 favored by the recommender system ($\alpha = 0.5$).

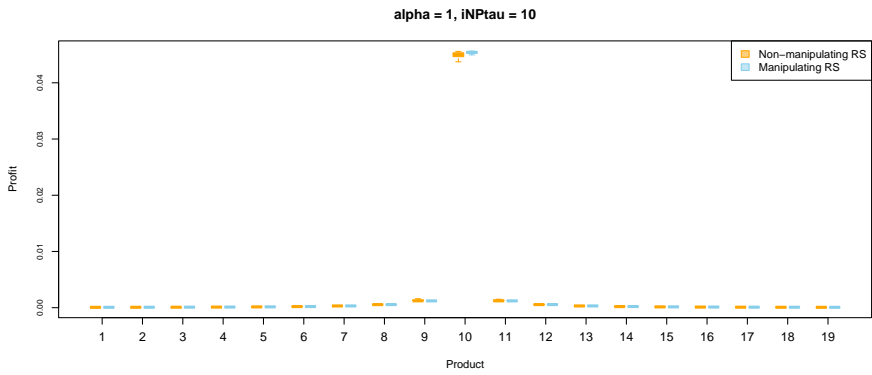


FIGURE 19.—Impact of 50% manipulation on profits for all products with product 10 favored by the recommender system ($\alpha = 1$).

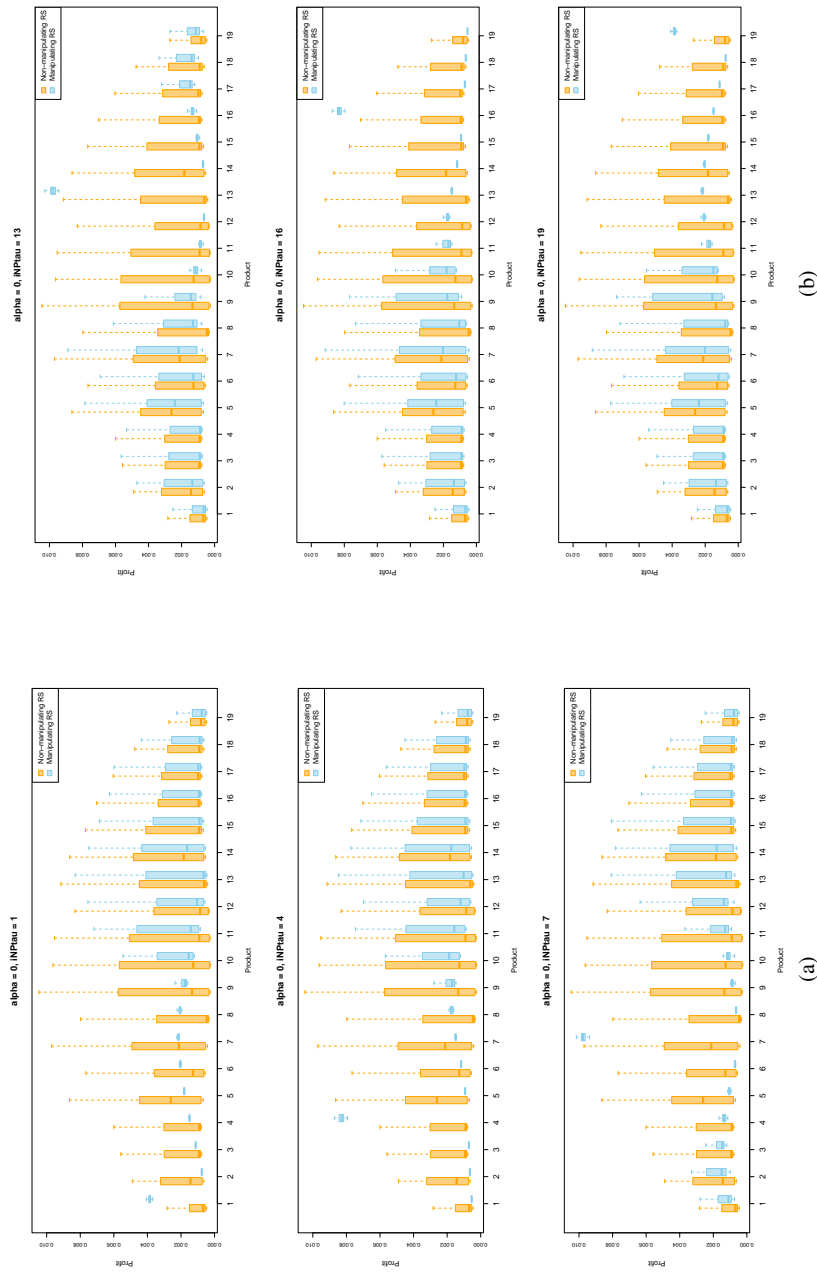


FIGURE 20.—Effect of manipulation on equilibrium profits, when the favored product differs from the central one. ‘INPtau’ is the favored product’s identity. Our labeling convention designates products based on their location, with products 1 and 19 as peripheral and product 10 as central. Key this figure: parameters $\alpha = 0$, and manipulation frequency maintained at 50%.

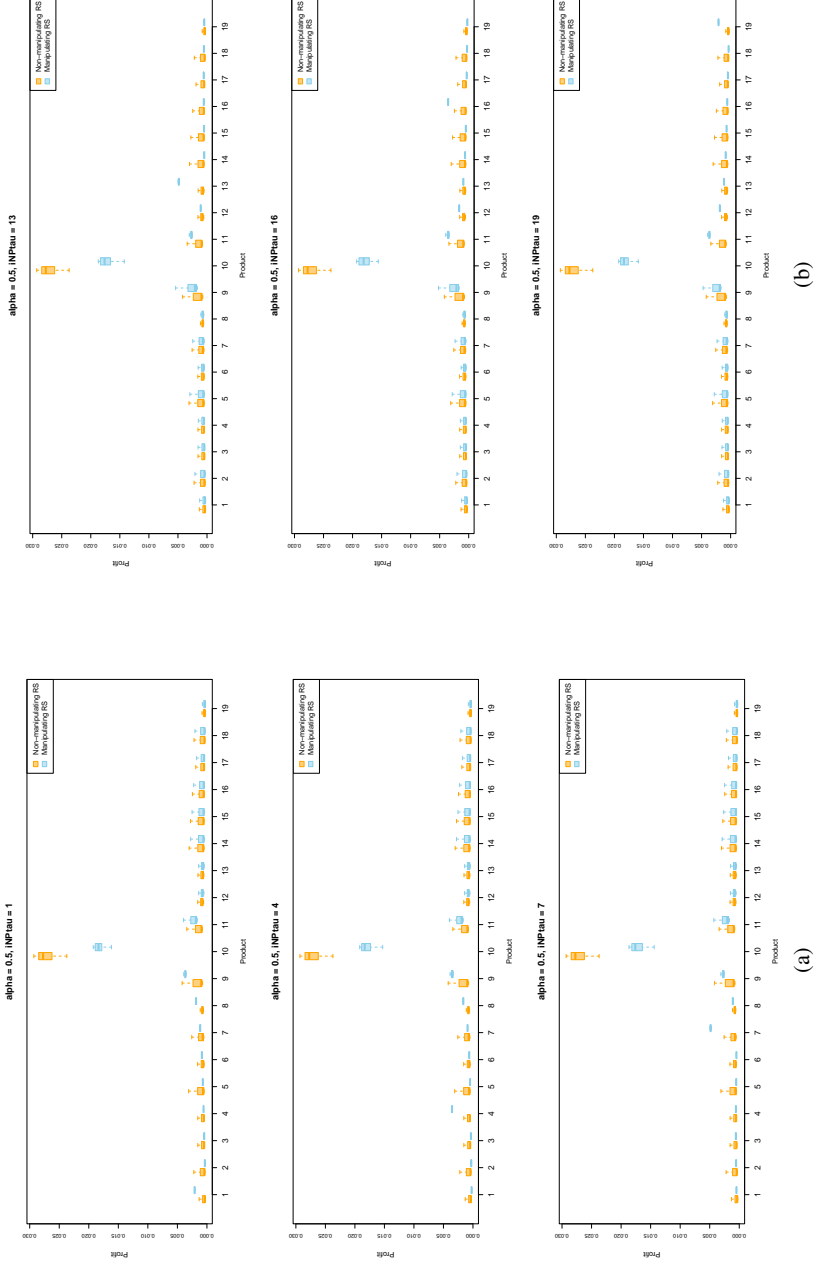


FIGURE 21.—Effect of manipulation on equilibrium profits, when the favored product differs from the central one. ‘INPtau’ is the favored product’s identity. Our labeling convention designates products based on their location, with products 1 and 19 as peripheral and product 10 as central. Key this figure: parameters $\alpha = 0.5$, and manipulation frequency maintained at 50%.

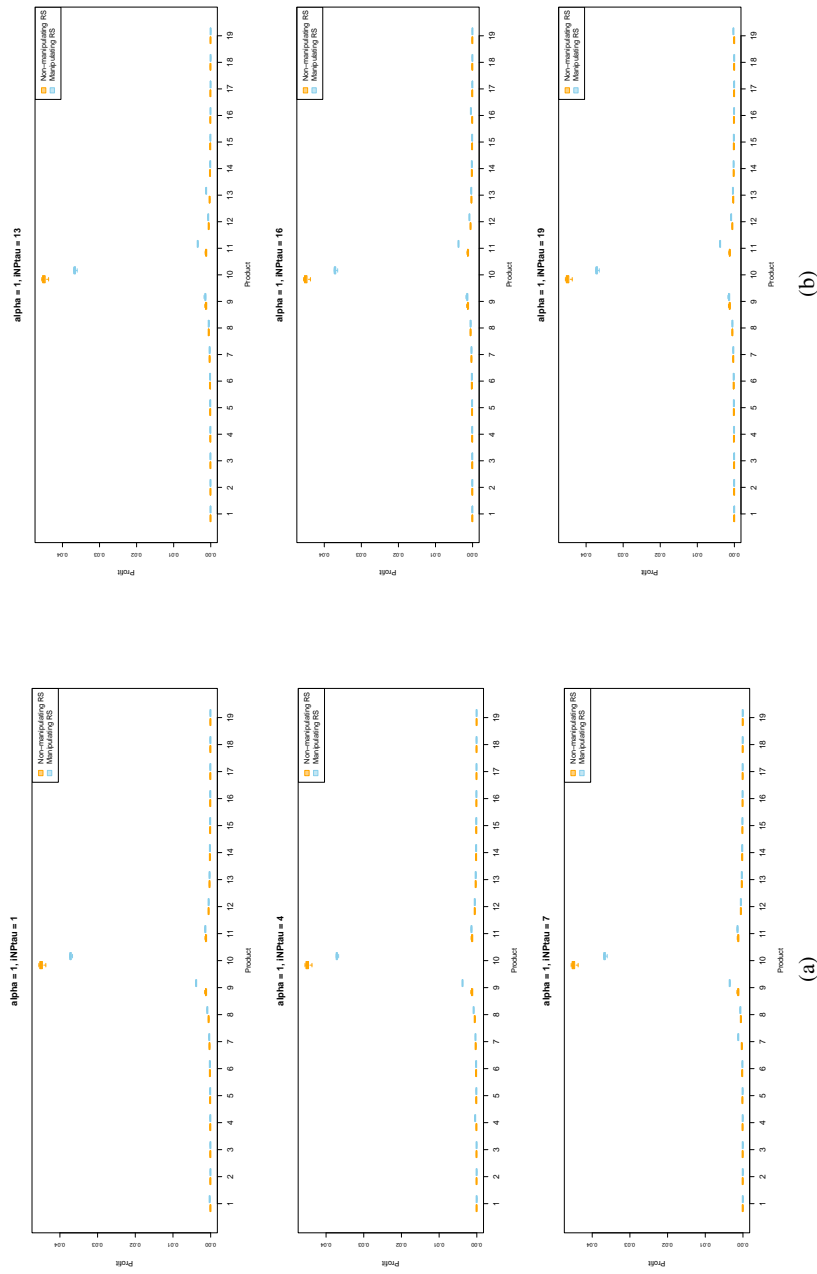


FIGURE 22.—Effect of manipulation on equilibrium profits, when the favored product differs from the central one. ‘INPtau’ is the favored product’s identity. Our labeling convention designates products based on their location, with products 1 and 19 as peripheral and product 10 as central. Key this figure: parameters $\alpha = 1$, and manipulation frequency maintained at 50%.

9.3. *Consumer surplus*

In Figure 23, consumer surplus decreases when recommendations are manipulated relative to the case where recommendations are sincere. This occurs independently of which product is favored with the manipulation and the level of α .

Note that when α is large, the scope for manipulation is quite limited even with a 50% manipulation rate. This is because in this case, most consumers are already directed towards the central product, regardless of manipulation. As a result, the impact of manipulation is almost imperceptible when α is close to 1.

This change is again driven by three factors: the quality of matching, search costs, and prices. Manipulation negatively affects the quality of matching, leading to an increase in search costs for consumers. However, it also causes a substantial decrease in prices which mitigates its negative impact on consumer surplus. The scenario in which Product 10 is favored and α is set to 1 holds a distinct significance, as manipulation enhances matching, leading to improved matches and consequently, an increase in surplus. All these figures confirm the results discussed in the main text.

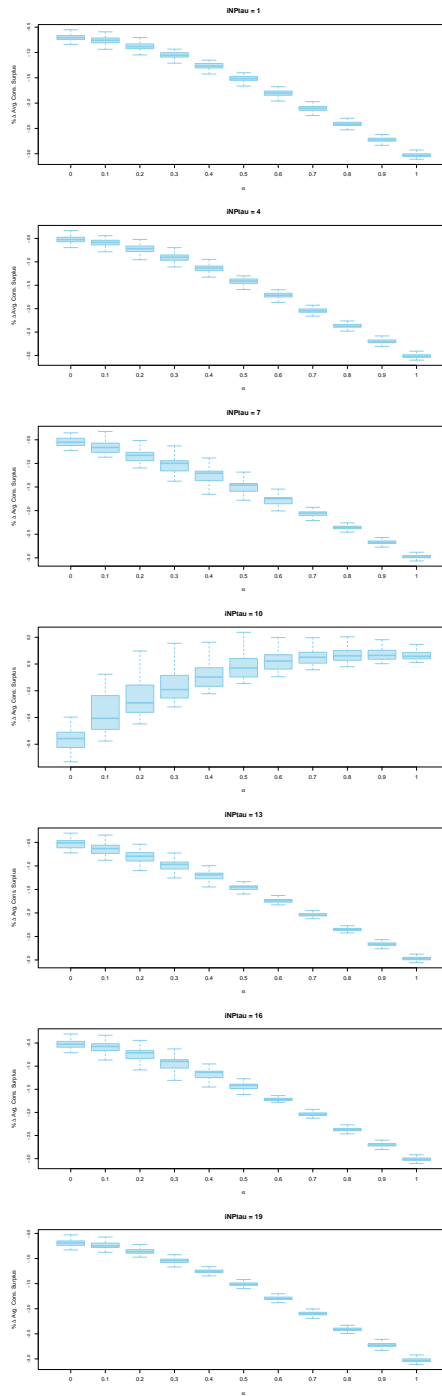


FIGURE 23.—Percentage difference in the average consumer surplus between the cases of manipulated and genuine recommendations with a rate of manipulation of 50%, for different manipulated products (labeled INPtau) and different values of α .