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"European Stock Market Dynamics Before and After the Introduction of the Euro"

by

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#### **European Stock Market Dynamics Before and After the Introduction of the Euro**

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#### Abstract

This paper addresses the following questions: Are the major European stock markets more integrated after the introduction of the Euro? How much of the change in the stock indices in different European countries can be attributed to innovations in other markets? How fast are events occurring in one European market transmitted to other markets? Vector Auto Regression models, impulses responses and variance decomposition are used to ascertain the stock market dynamics before and after the introduction of the Euro. The paper presents evidence of further integration of the European stock markets after the introduction of the Euro.

**JEL Classification:** F, G, C1, C3, C5, E44.

**Key Words:** Euro; Vector Auto Regression Models; Co-movements of Stock Markets; Impulse Response; Variance Decomposition.

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#### 1. Introduction

Many barriers to the free flow of goods, services, labor and capital between countries have been removed in the last several decades and national economies have become more open. This trend has been more pronounced in the fourteen Euro-zone countries in Europe, where a monetary union was implemented on January 1, 1999, with the Euro as the common currency. As a result, companies and investors in the Euro-zone can now make their investment decisions without worrying about exchange-rate risk. European stock markets have also experienced deregulation and integration. Many researchers have examined the linkages among national stock markets but the studies on the effect of the Euro are just beginning. The importance of the subject is clear from a recent keynote address by the president of the European Central Bank (ECB):

I believe that the key question for us – public authorities as well as market participants – is how we can contribute to the further integration of financial markets in Europe.[...] The potential gains from monetary union will only be fully realized if remaining barriers to integration of European financial markets are effectively removed. There is considerable evidence that wholesale markets are now much more integrated than before. But integration in securities markets needs to proceed further. Without an integrated European securities market the outcome of the entire process of financial market integration risks falling short of expectations.

> Keynote speech by ECB President Jean-Claude Trichet at Deutsche Börse's New Year's Reception 2004,

In a recent study, Bodart and Reding (1999) used conditional volatility models to show that, under the different stages of the European Monetary System, an increase of exchange rate volatility was associated with a decline in the correlation of national bond markets and an exchange rate peg was associated with a reduction of bond price volatility. However, these authors found only weak evidence on the interaction between exchange rate regime and equity market behavior.

This paper uses Vector Autoregressive approach to examine the changes in comovement among the stock market indices of the major Euro-zone countries: France, Germany, Netherlands, Italy, and Spain. Daily returns are analyzed for the period between January 1,1990 and May 31, 2003. Due to the dominance of the United Kingdom in the European stock market (see Friedman and Shachmurove, 1997), it is included in the analysis even though it is not part of the Euro-zone.

The key issue examined is how much of the change in the stock indices in different countries can be attributed to innovations in other markets, and how fast events occurring in one market are transmitted to other markets. Furthermore, the data is divided into two subperiods to investigate whether the integration of the stock markets increased after the adoption of the Euro. This paper is organized as follows: the next section discusses the reasons for co-movements among stock markets and why these linkages are thought to have become stronger in the recent years. Then a short survey of earlier studies of relationships among stock markets is presented in Section 3. Section 4 describes the framework of the analysis, including the data, the Vector Autoregressive (VAR) Model; unit roots tests and tests for determination of the lag length used in the VAR model. The empirical results are presented in Section 5, followed by the conclusions in Section 6.

#### 2. Reasons for co-movements among national stock markets

International stock prices are correlated for many reasons. First, the different stock markets may be influenced by the same macroeconomic variables, such as trade linkages between countries or booms and recessions in one country spilling over to other countries. For instance, the rise in an interest rate of one country, caused by high inflation, would lead to immediate fluctuations or interest rate movements in another country. The stock market returns in these two markets would be affected by this potential rise in the interest rate, causing stock prices to fall due to two well-known facts. An increase in interest rate makes it more attractive for investors to move their money away from stocks to other financial instruments such as bonds. In addition, the firms would face higher financial costs on their debt, which leads to a reduced cash flow (Durre and Giot, 2005).

Second, improved communication technology and the internet increased the speed of dissemination of news across the globe. Another contributing factor to markets' comovement is the higher degree of cooperation among national governments in recent years and the removal of trade barriers which prevented the flow of goods, services, and capital. This internationalization process has been evident in Europe, where the economic and financial structures have undergone extensive changes in the recent years. There has been a rapid development of the financial markets, which has been reinforced since the introduction of the Euro on January 1, 1999. The process of creating and the introduction of the Economic and Monetary Union have contributed to further economic integration and convergence of the economies inside the European Union (Noyer, 2000). First, within the Economic and Monetary Union there is a high degree of labor, goods, and capital mobility. Second, there is a common goal to achieve price stability. Third, the introduction of the Euro eliminated exchange rate risk inside the Union. Fourth, the macroeconomic policies in the participating countries are coordinated. Finally, the countries have a common monetary policy implemented by the European Central Bank (Apergis and Demopoulos, 1996). These factors are assumed to lead to a higher degree of co-movements among the stock market returns in Euro-zone countries. This assumption is examined in the paper.

#### **3.** Previous studies on the interdependence between stock markets

Since the work of Grubel in 1968, which pointed out the benefits of international diversification, there have been numerous studies examining the relationships among national stock markets. Earlier studies by Granger and Morgenstein (1970), Levy and Sarnet (1970), Grubel and Fadner (1971), Agmon (1972) and Ripley (1973) found little or no correlation between national stock markets, based on weekly or monthly data from the 1960's and 1970's. These studies used simple correlation and regression methods. The main conclusions that are found in these papers are (1) national stock markets are segmented and (2) risk reduction through international portfolio diversification is possible. The low degrees of co-movements between the stock markets are explained by barriers of international capital flows, different policies, higher taxes and transaction costs on international capital investments, and a low degree of information about foreign securities.

More recent research utilizes daily stock markets returns. Eun and Shim (1989) investigated the international transmission of stock markets movements using a Vector Autoregressive system, for the period 1980-1985. They found a substantial degree of interdependence among the nine stock markets in their study. Events occurring in the US stock market are quickly transmitted to stock markets in other countries. However, no individual stock market has substantial influence on the US stock market.

Gjerde and Sættem (1995) also use a Vector Autoregressive model to investigate the dynamic interactions among stock markets. They analyze the period between 1983 and 1994 by using ten stock market indices. The dynamic interactions among the stock markets are found to be larger than reported by Eun and Shim (1989). The study also found the US stock market to have significant influence on most other markets with the exception of Italy. On the other hand, the stock markets in Europe did not appear to have substantial influence on the world or any of the largest stock markets, such as New York and Tokyo. Furthermore, they found a rapid international transmission mechanism between the different stock markets. Most of the signals from one stock market can influence other markets within the same day; taking into account that stock markets operate in different time zones.

Friedman and Shachmurove (1997) also use a VAR-model in their research. They focus solely on European markets for the period January 1988 to December 1994. Their study finds a high degree of interdependence between the larger stock markets in Europe. However, the smaller markets prove to be more independent from other market fluctuations. None of the stock markets are found to be completely unaffected by innovations in other markets. The UK stock market proved to be the leading market in Europe during this time period, as opposed to the smaller markets, which seem to have no significant influence on the other markets.

The studies of Becker, Finnerty, and Gupta (1990) and Hassan and Naka (1996) also pointed out the leading role of the US stock market, using correlation analysis and a vector error correction model, respectively. Several researchers have analyzed if the co-movements between stock markets have become stronger after the October 1987 stock market crash. Jeon and Von Furstenberg (1990) analyzed the linkages among the stock markets in Japan, Germany, England, and USA, with the use of a VAR model for the period 1986-1988. By dividing this period into two parts, before and after the stock market crash, they state that for the period before the crash, the stock market returns can be explained based on innovations in the US stock market and from the history of each market. After the crash, the stock market returns were better explained based on previous changes in the foreign stock markets as opposed to their own history, with the exception of Japan. Using causality and co-integration tests, Malliaris and Urrutia (1992) and Arshanapalli and Doukas (1993) also show that the degree of interdependence among stock markets has increased significantly after the 1987 stock market crash.

Malkamäki, et al, (1993) focus on the co-movements among the Scandinavian stock markets. They analyze the stock markets in Sweden, Norway, Denmark, and Finland for the period between February 1988 and April 1990, using Granger causality tests. The Swedish stock market has been found to lead the group, whereas the rest show no significant influence over other markets. They also determined that returns on the world stock market, using the world stock index, have a significant leading effect on the Scandinavian stock market returns.

Kana (1998) analizes the linkages among the US stock market and six major European stock markets for the period January 1983 to November 1996. The results from this research differ from the findings in the other analyses above. Using three different methodologies to test for co-integration, Kana discovered that the US stock market does not have pair wise co-integration with any of the European markets. These results imply that there are potential benefits from diversifying in US stocks as well as stocks in European markets. Short-run and long-term links between European and US stock markets are analyzed by Gerrits and Yüce (1999), using a vector error correction model for the period between March 1990 and October 1994. The US market has a long-term influence on the European markets but this is not true in reverse. Also, in a short-term perspective the US market shows a substantial influence on all other markets. The European stock markets are found to move together. Compared to the studies from the 1960's and 1970's, the more recent studies report a substantial amount of interdependence among national stock markets. This phenomenon can be explained as being the result of the removal of barriers for foreign investment, improvement in information technology, and increased co-operation and trade between countries, along with other reasons.

## 4. Methodology

This study aims to examine changes in co-movements among the European stock markets after the introduction of the Euro. By the use of the Vector Autoregressive approach, stock market indices of the major Euro-zone countries, such as France, Germany, Netherlands, Italy, and Spain, are analyzed based on daily returns for the period between January 1, 1990 and May 31, 2003. Due to the dominance of the United Kingdom in the European stock market, it is also included in the analysis even though it is not part of the Euro-zone.

## 4.1 Data

The data for this study comprise time series of daily stock market indices for six major European stock exchanges: Britain, France, Germany, Italy, Netherlands, and Spain. Morgan Stanley Capital International (MSCI) computed the indices for 1/1/1990 to 05/31/2003. For the methodology of constructing the indices, see <u>http://www.msci.com/method/index2.html</u>.<sup>1</sup> The indices are expressed in German Marks for the pre-Euro period and in Euros for the post Euro period. For each country, daily returns,  $r_t$ , are computed as the first differences of the natural logarithms of  $P_t$ , the daily close values of the indices,  $r_t = (lnP_t-lnP_{t-1})$ .

#### 4.2 Correlation Analysis

Table 1 shows the correlations among the daily returns for the entire sample period and for the two sub periods. Note that the EC stock markets are highly correlated as is evidenced by the correlations between the stock markets of France and Netherlands (0.80), Britain and Netherlands (0.75), France and Spain (0.75), Britain and France (0.74), Germany and France (0.74), and Germany and Netherlands (0.74). Thus, it is apparent that geographical proximity matters. It is also worth noting that the Italian stock market has the lowest correlations with other EC stock markets.

Insert TABLE 1 about here

There is a striking difference in correlations between the two sub periods. For example, the correlation between UK and France increased from 0.70 to 0.79 in the second sub-period. The correlation between UK and Germany increased from 0.55 to 0.73, and the correlation between France and Spain increased from 0.69 to 0.83 in the same period.

Each country's series of daily returns (expresses as  $r_t = (lnP_t-lnP_{t-1}))$  was tested for the presence of a unit root using three alternative tests suggested by Dickey and Fuller [1979], and Phillips and Perron [1998]. All these tests reject the assumption of a unit root for all time series considered, implying that the relationships among the various variables analyzed below are not spurious.

#### 4.3 Model Choice

The analysis uses a Vector Autoregression Approach, which is described in Appendix A. Sims (1980) criticized the simultaneous literature for the ad hoc restrictions needed for identification and for the ad hoc classification of exogenous and endogenous variables in the system. The main differences from the traditional structural (Cowles Commission) method of constructing such econometric models are (Charemza and Deadman, 1997):

- There is not a prior endogenous division of variables. There are no stock markets that are exogenous. It is not the case that the stock market in Germany influences the stock market in UK, and not the other way around. In a Vector Autoregressive (VAR) Model one does not have to indicate which stock market is exogenous and which is endogenous. However, this is necessary in the structural single-equation or multi equation co-integration models.
- 2. No zero restrictions are imposed. Using the VAR Model one tests how the stock markets influence one another simultaneously. In a simple co-integration approach, one analyzes the direct effect of a one-unit rise in the German stock market on the UK stock market, when all other variables are held constant. This restriction is in fact extremely unrealistic.
  - 3. There is no strict (and prior to modeling) economic theory within which the model is grounded. This follows from the previous points. One does not have to decide which stock market influences another, and no restrictions are imposed. A VAR Model requires minimal theoretical demands on the model structure.

With a VAR Model, one only needs to make two specifications (Pindyck and

Rubinfeld, 1998, p. 400):

- (1) The variables (endogenous and exogenous) that are believed to interact should be included as a part of the economic system one is trying to model.
- (2) The largest number of lags needed to capture most of the effects that the variables have on each other.

The two most common methods used in the recent analysis of linkages among stock markets are the Vector Autoregressive Model and a Vector Error Correction Model. As the aim of this paper is to analyze linkages among European stock markets, we chose to use the VAR Model because it was found to be more appropriate for studying dynamic inter-linkages of stock markets. Since the main concern of investors is the stock market returns and not some arbitrary defined index levels, we believe that the relevant co-integration is between returns and not the levels (see for example Friedman and Shachmurove, 1997 and Gjerde and Sættem, 1995).

One weakness of the VAR Model is the importance of ordering the variables. It is important which stock market is called y<sub>1</sub>, which is called y<sub>2</sub>, and so on. The order of the variables has to be specified by the analyst, since no statistical methods exist that can determine the ordering of the variables. In an impulse response analysis the first variable must therefore be the only one with a potential immediate impact on all other variables. The second could influence the remaining variables, but it would have no potential immediate impact on the first variable. Another problem is the potential for the model to be incomplete. When important variables are omitted from the system, this could have an impact on the results, since it makes the impulse responses less valuable for structural interpretations.

## 4.4 Determination of the Lag-Length

An important part of the analysis is to determine the appropriate lag structure in the VAR Model. Standard recommendations for the selection of the appropriate lag length is to choose the number of lags to be long enough to ensure that the residuals are white noise, but not too long since the estimates can become imprecise. The lag length is therefore often selected somewhat arbitrarily. The number of lags was chosen based on three tests: The Likelihood Ratio tests (Sims, 1972), the Information Criteria suggested by Akaike (1973) and by Schwarz (1968). While the Akaike and Schwarz tests indicated that as few as 2 lagged daily returns may be sufficient, the Sims test suggested that 15 lags are needed. A lag length of 15

ensures that all the dynamics in the data is captured and is used in this analysis. (Eun and Shim [5] and Friedman and Shachmurove (1997) also used 15 lags.)

#### 5. Empirical Results

Table 2 presents the correlation matrix of residual returns. These residuals are the part of returns not explained by past returns of all six stock markets. The larger stock markets in Europe have high correlation values, a fact which will affect the inferences from the variance decomposition below. Like the correlation of returns, the correlation of residuals also increased following the introduction of the Euro.

Insert TABLE 2 about here

Table 3 presents the results of the forecast error variance decomposition for 5-day, 10-day, and 15-day-ahead horizons for three periods. The first part consists of the entire sample, namely from January 1, 1990 until the end of the sample, May 31, 2003. The remaining parts are two sub-periods; Period I is before the introduction of the Euro, from January 1, 1990 until December 31, 1998 and Period II is after the introduction of the Euro, from January 1, 1999 until May 31, 2003. The Cholesky ordering reported here are as follows: UK, Germany, France, Holland, Italy, and Spain where the stock markets were ranked by market capitalization.

Insert TABLE 3 about here

For all periods of analysis, within a time horizon of 5, 10, and 15 days, Holland is the most open market, being influenced by about 72 percent from all other markets excluding its own market. The next most open market throughout the period of the study is France (69

percent) and then Spain (about 63 percent) and Italy (48 percent). The leading market in Europe is the UK, where about 40 percent of a one standard deviation shock to Germany and to Spain is explained by the UK stock market returns. The UK market explains about 54 percent of the French and Netherlands stock market returns. As Friedman and Shachmurove (1997) have found, the ordering between France and Germany may influence the results reported here, increasing the openness of Germany's stock market to outside innovations and decreasing the openness of the French stock market.

Comparing the variance decompositions for the first and the second periods, it is clear that the European stock markets have become more integrated after the introduction of the Euro. For example, whereas the UK stock market was influenced only by 4 percent of other markets in the first period, this number more than doubled in the second period. The results for all other markets in the study are indeed remarkable. Germany is now being influenced by 57 percent; compared with only 35 percent in the first period. The influence other markets now have on France has increased from 60 percent to 80 percent. The Netherlands market, which was already an open market, was influenced by 65 percent in the pre-Euro period and is now being influenced as much as the French stock market. Moreover, after the introduction of the Euro, the influence on the relatively isolated Italian market increased to 76 percent, after being preceded by only 38 percent in the first period. Another remarkable change occurred in the Spanish market; it is now explained by the influence of the other markets by 73 percent as compared to 58 percent before the introduction of the Euro.

Table 4 and Figure 1 present the simulation results of the accumulated responses of a onestandard-deviation shock, using Cholesky ordering based on Monte Carlo with 100 repetitions. The ordering is again as in Table 3. Table 4 presents these accumulated responses for the two sub-periods, Period I and Period II, and only for the first three days after the introduction of the shocks. For example, the effect of the UK market on the German market increased from 0.72 percent before the introduction of the Euro to 1.25 percent after the Euro's, introduction, a 74 percent increase. The accumulated effect of the German market on the French market rose from 0.21 percent to 0.80 percent. In general, one can see the same pattern emerging; more interdependence and a higher degree of inter-linkages among the developing stock markets of Europe.

Insert TABLE 4 about here

Insert Figure 1 about here

#### 6. Conclusions

Using a Vector Autoregressive Model this paper analyzes causal relations and dynamic interactions between major Euro zone stock markets. The data covers a period with large changes in the European economies, with the process towards and the introduction of the Economic and Monetary Union. The findings in this study indicate that the comovements of the European stock markets have increased after the introduction of the Euro. This is especially pronounced for the stock markets in Germany, France, Holland, Italy and Spain, the five Euro zone countries in the analysis. Also compared to previous studies, the results indicate substantial increased international financial integration. This implies that the benefits of international diversification within the Euro zone stock markets have decreased considerably over the recent years. Some interesting results are found for single countries. For instance, the Italian stock market is found to be linked to the other European markets to a much higher extent than was found in previous studies. Finally, the time paths of impulse responses to a shock in the stock market in the UK, the only country in the study that kept its national currency, shows a rapid transmission of information between stock markets in Europe. This supports the hypothesis of international stock markets efficiency.

#### Footnote

1. To construct an MSCI Country Index, every listed security in the market is identified. Securities are free float adjusted, classified in accordance with the Global Industry Classification Standard and screened by size and liquidity. MSCI then constructs its indices by targeting for index inclusion 85% of the free float adjusted market capitalization in each industry group, within each country. By targeting 85% of each industry group, the MSCI Country Index captures 85% of the total country market capitalization while it accurately reflects the economic diversity of the market.

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#### APPENDIX A.

#### THE VAR MODEL

The VAR model assumes that each variable depends on its own past values and on the past values of all other variables in the system of equations. The model can be expressed as

(A.1) 
$$Y_t = X_t \bullet \boldsymbol{b} + \sum_{s=1}^{L} A_s \bullet Y_{t-s} + U_t$$
  
(A.2)  $E[U_t \bullet U_{t'}] = \Psi$ 

where  $Y_t$  is an n×1 vector of daily stock market returns,  $X_t \cdot \beta$  is the deterministic component of  $Y_t$ . In the present application  $X_t$  is a vector of ones. The term  $U_t$  is an n×1 vector of serially uncorrelated errors,  $A_s$  is an n×n matrix of coefficients and L is the number of lags. The moving average representation (MAR) of the VAR model can be written as

$$(A.3) \quad Y_t = X_t \bullet \boldsymbol{b} + \sum_{s=0}^{\infty} B_s \bullet E_{t-s}$$

where,  $E_{t-s}$  for s=0,..., $\infty$  is an n-variate white noise process, and  $E_t$  and  $E_s$  are uncorrelated for t + s, (Sims [1972).

There are many equivalent representations for this model. For any non-singular matrix G, the matrix of coefficients  $B_s$  can be replaced by  $B_s \cdot G$  and E by  $G^1 \cdot E$ . A particular version is obtained by choosing some normalization.

If  $B_0$  is normalized to be the identity matrix, each component of  $E_t$  is the error that results from the one step ahead forecast of the corresponding components of  $Y_t$ . These are the non-orthogonal innovations in the components of Y because, in general, the covariance matrix  $\Phi = E[E_t \cdot E_t']$  is not diagonal.

It is more useful to look at the moving average representation of the system with orthogonalized innovations. If any matrix G is constructed to satisfy

$$(A.4) \quad G^{-1} \bullet \Phi \bullet \quad G^{-1} = h$$

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then the new innovations  $v_t = E_t \cdot G^{-1}$  satisfy

$$(A.5) \quad E[v(t) \bullet v(t)'] = I$$

These orthogonalized innovations have the important property that they are uncorrelated across time and across equations. Such a matrix G can be any solution which satisfies the condition that  $GG' = \Phi$ . The problem is that there are many such factorizations of a positive definite matrix  $\Phi$ .

The literature on time-series suggests a number of ways to accomplish the factorization of  $\Phi$ . Some techniques are based on the Choleski factorization, where G is restricted to be a lower triangular matrix. Other techniques are based on orthogonalization using the eigenvalues. Sims (1980) suggested imposing restrictions on the  $\Phi$  matrix by constraining it to be a lower triangular matrix. In general, the moving average model (A.4) is diagonalized as follows:

$$(A.6) \quad BU(t) = V(t)$$

and

$$(A.7) \quad E[V(t) \bullet V(t)'] = D$$

where D is a diagonal matrix. The model can be estimated by minimizing the log likelihood function with respect to the free parameters in the matrices, A and D in equation (A.8).

(A.8) 
$$-2\log_{1}A_{1}^{\prime} + \log_{1}D_{1}^{\prime} + \operatorname{trace}(D^{-1}\cdot A \cdot S \cdot A^{\prime})$$

where S is the sample covariance matrix of residuals, and A is the coefficients matrix of (A.1).

Table 1	: Correla	tion Matri	x 1/1/1990	- 5/31/2003	3				
	UK	FRA	GER	HOL	ITA	SPA			
UK	1.00	0.74	0.63	0.75	0.58	0.64			
FRA	0.74	1.00	0.74	0.80	0.65	0.75			
GER	0.63	0.74	1.00	0.74	0.60	0.66			
HOL	0.75	0.80	0.74	1.00	0.62	0.69			
ITA	0.58	0.65	0.60	0.62	1.00	0.64			
SPA	0.64	0.75	0.66	0.69	0.64	1.00			
Table 1: Correlation Matrix 1/1/1990 – 12/31/1998									
	UK	FRA	GER	HOL	ITA	SPA			
UK	1.00	0.70	0.55	0.72	0.50	0.61			
FRA	0.70	1.00	0.64	0.71	0.55	0.69			
GER	0.55	0.64	1.00	0.67	0.49	0.59			
HOL	0.72	0.71	0.67	1.00	0.52	0.63			
ITA	0.50	0.55	0.49	0.52	1.00	0.56			
SPA	0.61	0.69	0.59	0.63	0.56	1.00			
Table 1	: Correla	tion Matri	x 1/1/1999	- 5/31/2003	3				
	UK	FRA	GER	HOL	ITA	SPA			
UK	1.00	0.79	0.73	0.80	0.72	0.69			
FRA	0.79	1.00	0.84	0.88	0.84	0.83			
GER	0.73	0.84	1.00	0.80	0.79	0.75			
HOL	0.80	0.88	0.80	1.00	0.80	0.78			
ITA	0.72	0.84	0.79	0.80	1.00	0.80			
SPA	0.69	0.83	0.75	0.78	0.80	1.00			

 Table 1:
 Correlation Matrix for sub-periods

Table 2: Correlation Matrix of Residuals 12/31/1990 - 5/31/2003										
	UK	FRA	GER	HOL	ITA	SPA				
UK	1.000	0.738	0.636	0.750	0.577	0.640				
FRA	0.738	1.000	0.751	0.793	0.653	0.749				
GER	0.636	0.751	1.000	0.747	0.597	0.665				
HOL	0.750	0.793	0.747	1.000	0.615	0.690				
ITA	0.577	0.653	0.597	0.615	1.000	0.636				
SPA	0.640	0.749	0.665	0.690	0.636	1.000				

 Table 2:
 Correlation Matrices of Residual Returns

Table 2: Correlation Matrix of Residuals 12/31/1990 - 12/31/1998										
	UK	FRA	GER	HOL	ITA	SPA				
UK	1.000	0.696	0.556	0.711	0.505	0.608				
FRA	0.696	1.000	0.651	0.713	0.551	0.695				
GER	0.556	0.651	1.000	0.675	0.487	0.596				
HOL	0.711	0.713	0.675	1.000	0.515	0.625				
ITA	0.505	0.551	0.487	0.515	1.000	0.549				
SPA	0.608	0.695	0.596	0.625	0.549	1.000				

Table 2: Correlation Matrix of Residuals 12/31/1998 - 5/31/2003										
	UK	FRA	GER	HOL	ITA	SPA				
UK	1.000	0.772	0.719	0.776	0.699	0.659				
FRA	0.772	1.000	0.845	0.853	0.838	0.811				
GER	0.719	0.845	1.000	0.792	0.774	0.737				
HOL	0.776	0.853	0.792	1.000	0.783	0.754				
ITA	0.699	0.838	0.774	0.783	1.000	0.786				
SPA	0.659	0.811	0.737	0.754	0.786	1.000				

Table 3	Table 3a: Markets Influences Explained during 01/01/1990 to 05/31/2003										
	Period	UK	GER	FRA	HOL	ITA	SPA	AOM*			
UK	5	98.99	0.25	0.09	0.42	0.12	0.13	1.01			
	10	97.35	0.80	0.33	0.67	0.33	0.51	2.65			
	15	96.49	0.98	0.44	0.91	0.44	0.74	3.51			
GER	5	39.70	58.96	0.97	0.09	0.11	0.19	41.04			
	10	39.72	58.03	1.25	0.40	0.22	0.38	41.97			
	15	39.42	57.77	1.41	0.55	0.32	0.53	42.23			
FRA	5	53.94	13.67	31.80	0.34	0.19	0.07	68.20			
	10	53.46	13.74	31.24	0.79	0.33	0.44	68.76			
	15	53.08	13.93	30.99	0.93	0.42	0.66	69.01			
HOL	5	55.73	12.42	3.86	27.81	0.13	0.04	72.19			
	10	55.15	12.49	4.06	27.50	0.32	0.48	72.50			
	15	54.57	12.56	4.13	27.56	0.37	0.81	72.44			
ГТА	5	32.99	8.88	4.52	0.85	52.73	0.03	47.27			
	10	32.83	8.91	4.67	1.37	52.13	0.10	47.87			
	15	32.58	9.18	4.66	1.39	51.66	0.52	48.34			
SPA	5	40.49	11.55	7.41	0.81	2.26	37.47	62.53			
	10	40.33	11.67	7.38	1.37	2.33	36.91	63.09			
	15	40.06	11.97	7.46	1.50	2.44	36.57	63.43			

Table 3:Results from the Variance Decomposition

# Table 3: continued

			Before In	troductio	n of Euro			
Table 3	3b: Market	ts Influen	ces Expla	nined for	period 01	/01/1990	to 12/31/	1998
	Period	UK	GER	FRA	HOL	ITA	SPA	AOM*
UK	5	98.15	0.84	0.09	0.60	0.11	0.21	1.85
	10	96.88	1.19	0.39	0.63	0.62	0.30	3.12
	15	95.92	1.31	0.45	0.78	0.73	0.80	4.08
GER	5	29.67	66.95	2.50	0.22	0.18	0.48	33.05
	10	29.47	65.95	3.16	0.42	0.44	0.55	34.05
	15	29.63	65.28	3.13	0.49	0.59	0.88	34.72
FRA	5	47.83	10.99	40.73	0.17	0.25	0.03	59.27
	10	47.32	11.14	40.34	0.29	0.48	0.43	59.66
	15	47.07	11.24	39.90	0.40	0.60	0.79	60.10
HOL	5	49.59	12.15	2.90	35.06	0.18	0.12	64.94
	10	49.09	12.16	3.34	34.70	0.52	0.19	65.30
	15	48.66	12.24	3.36	34.39	0.64	0.70	65.61
ITA	5	24.87	6.28	4.06	0.96	63.67	0.16	36.33
	10	25.08	6.24	4.18	1.48	62.74	0.28	37.26
	15	25.15	6.43	4.13	1.70	61.77	0.82	38.23
SPA	5	35.96	10.59	7.34	0.75	2.27	43.09	56.91
	10	35.65	10.78	7.35	1.36	2.42	42.44	57.56
	15	35.66	10.92	7.36	1.47	2.55	42.04	57.96

**Before Introduction of Euro** 

	-					=.00					
			After Int								
Table 3	Table 3c: Markets Influences Explained for period 12/31/1998 to 5/31/2003										
	Period	UK	GER	FRA	HOL	ITA	SPA	AOM*			
UK	5	98.10	0.63	0.21	0.50	0.43	0.13	1.90			
	10	93.53	1.90	1.24	1.31	0.55	1.47	6.47			
	15	91.22	2.40	1.72	2.03	0.81	1.82	8.78			
GER	5	52.37	46.39	0.15	0.28	0.62	0.19	53.61			
	10	52.37	43.85	1.42	0.59	0.74	1.02	56.15			
	15	50.65	43.20	2.26	0.93	1.03	1.93	56.80			
FRA	5	59.60	17.21	21.24	0.78	0.94	0.23	78.76			
	10	57.99	16.85	20.58	1.91	1.07	1.60	79.42			
	15	56.37	17.12	20.35	2.56	1.21	2.39	79.65			
HOL	5	60.34	12.19	5.80	20.50	0.72	0.46	79.50			
	10	58.51	12.13	6.29	19.95	1.11	2.01	80.05			
	15	56.58	12.44	6.54	20.31	1.21	2.92	79.69			
ITA	5	50.12	14.97	7.62	1.14	25.84	0.31	74.16			
	10	49.59	14.62	7.81	1.62	24.88	1.48	75.12			
	15	48.45	15.08	8.11	1.89	24.31	2.17	75.69			
SPA	5	45.19	14.32	8.20	1.45	2.73	28.11	71.89			
	10	44.37	14.25	8.98	2.27	2.84	27.29	72.71			
	15	43.16	14.55	9.41	2.78	2.87	27.23	72.77			

Cholesky Ordering: UK GER FRA HOL ITA SPA Standard Errors: Monte Carlo (100 repetitions)

\*AOM: All Other Markets. Denotes the percentage of forecast error variance of the market of the first column explained collectively by all the other markets.

Table		Introducti					
	Period	UK	GER	FRA	HOL	ITA	SPA
UK	1	1.11%	0.00%	0.00%	0.00%	0.00%	0.00%
	2	1.03%	-0.06%	-0.02%	-0.07%	0.01%	0.03%
	3	1.06%	-0.13%	-0.01%	-0.10%	0.04%	0.06%
GER	1	0.65%	0.98%	0.00%	0.00%	0.00%	0.00%
	2	0.72%	0.86%	0.19%	-0.02%	0.04%	0.07%
	3	0.72%	0.79%	0.21%	-0.06%	0.06%	0.08%
FRA	1	0.84%	0.38%	0.78%	0.00%	0.00%	0.00%
	2	0.79%	0.32%	0.80%	-0.05%	0.02%	0.00%
	3	0.76%	0.21%	0.85%	-0.05%	0.07%	-0.01%
HOL	1	0.75%	0.35%	0.18%	0.62%	0.00%	0.00%
	2	0.71%	0.29%	0.20%	0.55%	0.02%	0.03%
	3	0.71%	0.20%	0.21%	0.50%	0.05%	0.03%
ITA	1	0.78%	0.38%	0.29%	0.11%	1.24%	0.00%
	2	0.81%	0.40%	0.39%	0.01%	1.40%	0.04%
	3	0.79%	0.31%	0.42%	0.00%	1.44%	0.07%
SPA	1	0.83%	0.42%	0.36%	0.09%	0.19%	0.90%
	2	0.80%	0.41%	0.42%	0.03%	0.23%	1.00%
	3	0.77%	0.27%	0.45%	0.00%	0.29%	1.03%
Table 4		troduction					
	Period	UK	GER	FRA	HOL	ITA	SPA
UK	1	1.42%	0.00%	0.00%	0.00%	0.00%	0.00%
	2	1.37%	0.11%	0.00%	-0.04%	-0.01%	0.00%
	3	1.26%	0.13%	0.03%	-0.04%	0.07%	0.00%
GER	1	1.36%	1.28%	0.00%	0.00%	0.00%	0.00%
	2	1.28%	1.25%	0.03%	0.00%	0.01%	0.01%
	3	1.25%	1.29%	0.05%	0.00%	0.13%	0.05%
FRA	1	1.26%	0.65%	0.74%	0.00%	0.00%	0.00%
	2	1.27%	0.82%	0.63%	-0.05%	-0.01%	-0.06%
	3	1.19%	0.80%	0.60%	-0.05%	0.12%	-0.08%
HOL	1	1.28%	0.54%	0.38%	0.73%	0.00%	0.00%
	2	1.29%	0.71%	0.32%	0.63%	-0.02%	-0.08%
	3	1.21%	0.71%	0.30%	0.64%	0.11%	-0.11%
ITA	1	1.04%	0.56%	0.40%	0.11%	0.74%	0.00%
	2	1.01%	0.63%	0.38%	0.09%	0.68%	-0.05%
	3	0.97%	0.70%	0.36%	0.07%	0.75%	-0.07%
~							

1

2

3

1.08%

1.05%

0.94%

0.60%

0.72%

0.71%

0.45%

0.38%

0.39%

0.14%

0.12%

0.13%

0.25%

0.24%

0.32%

SPA

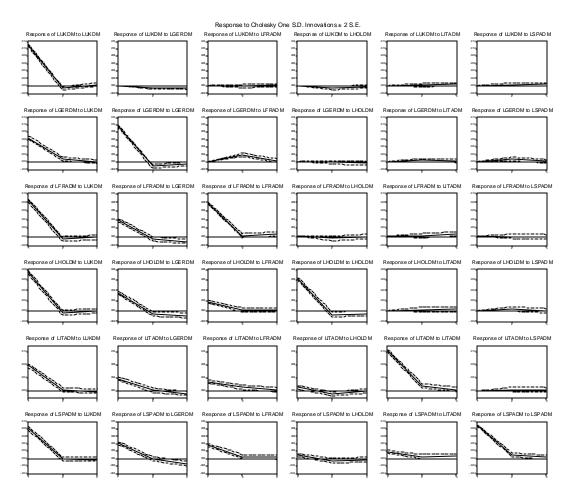
Table 4:Accumulated Responses of a One Standard Deviation shock using<br/>Cholesky Ordering based on Monte Carlo (100 repetitions) Simulations

0.85%

0.85%

0.82%

# Figure 1 A: GRAPH BEFORE: Response to Cholesky One S.D. Innovations ± 2 S.E.



# Figure 1B: GRAPH AFTER: Response to Cholesky One S.D. Innovations ± 2 S.E.

